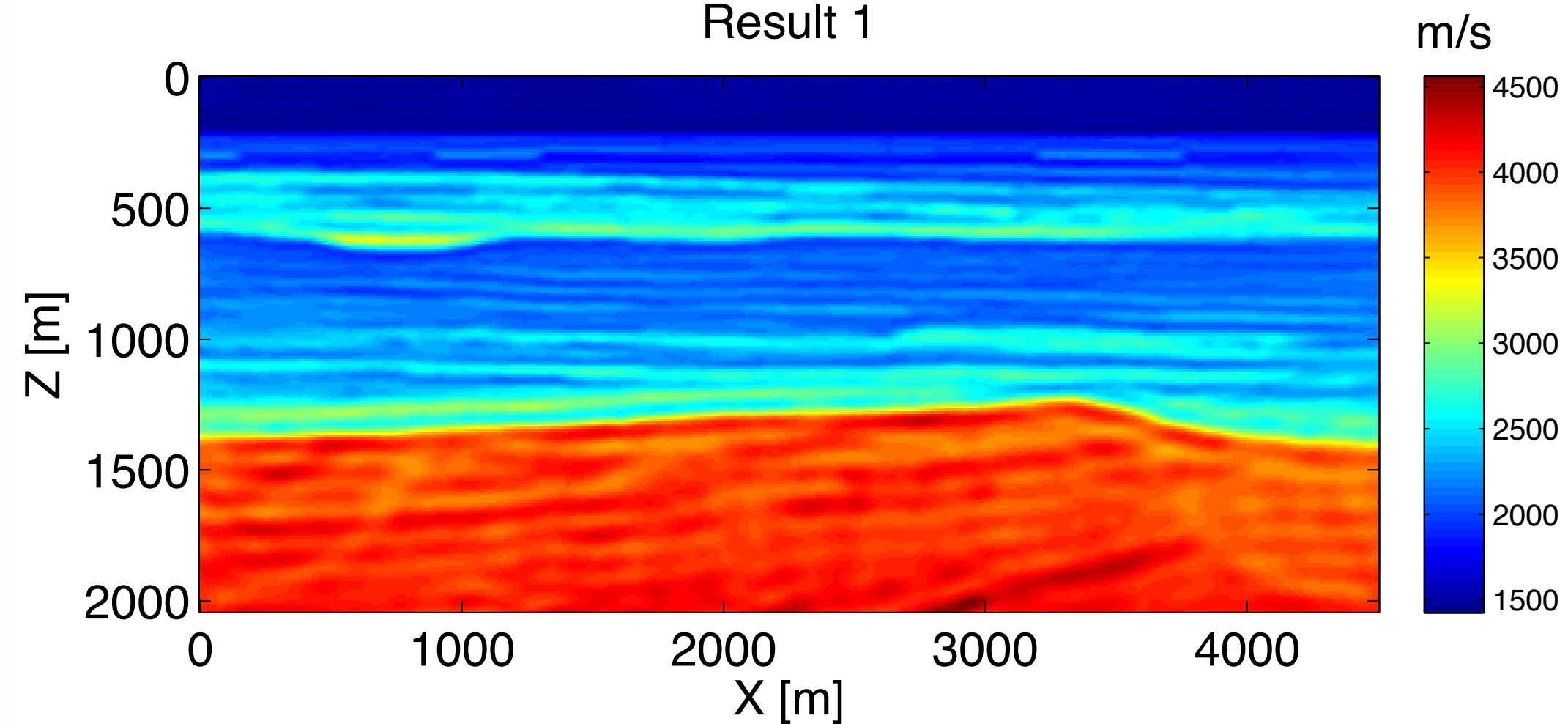


A stochastic quasi-Newton MCMC method for uncertainty quantification of full-waveform inversion

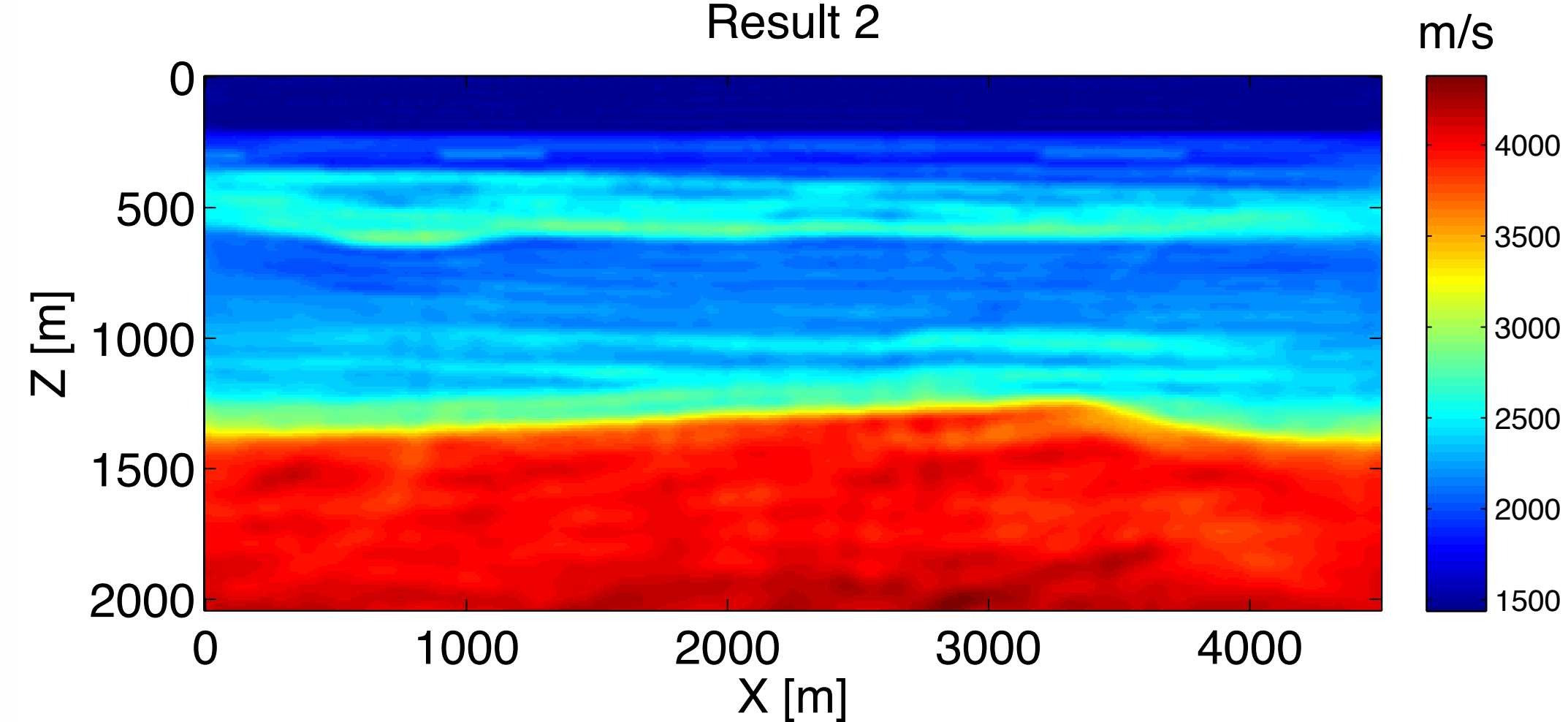
Zhilong Fang, Felix J. Herrmann and Chia Ying Lee

Motivation

Result 1



Result 2



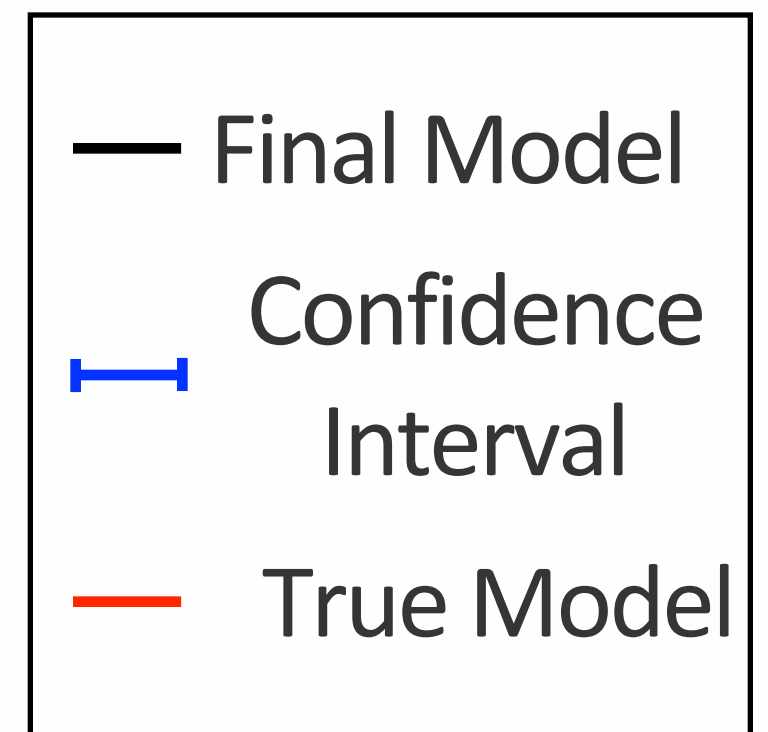
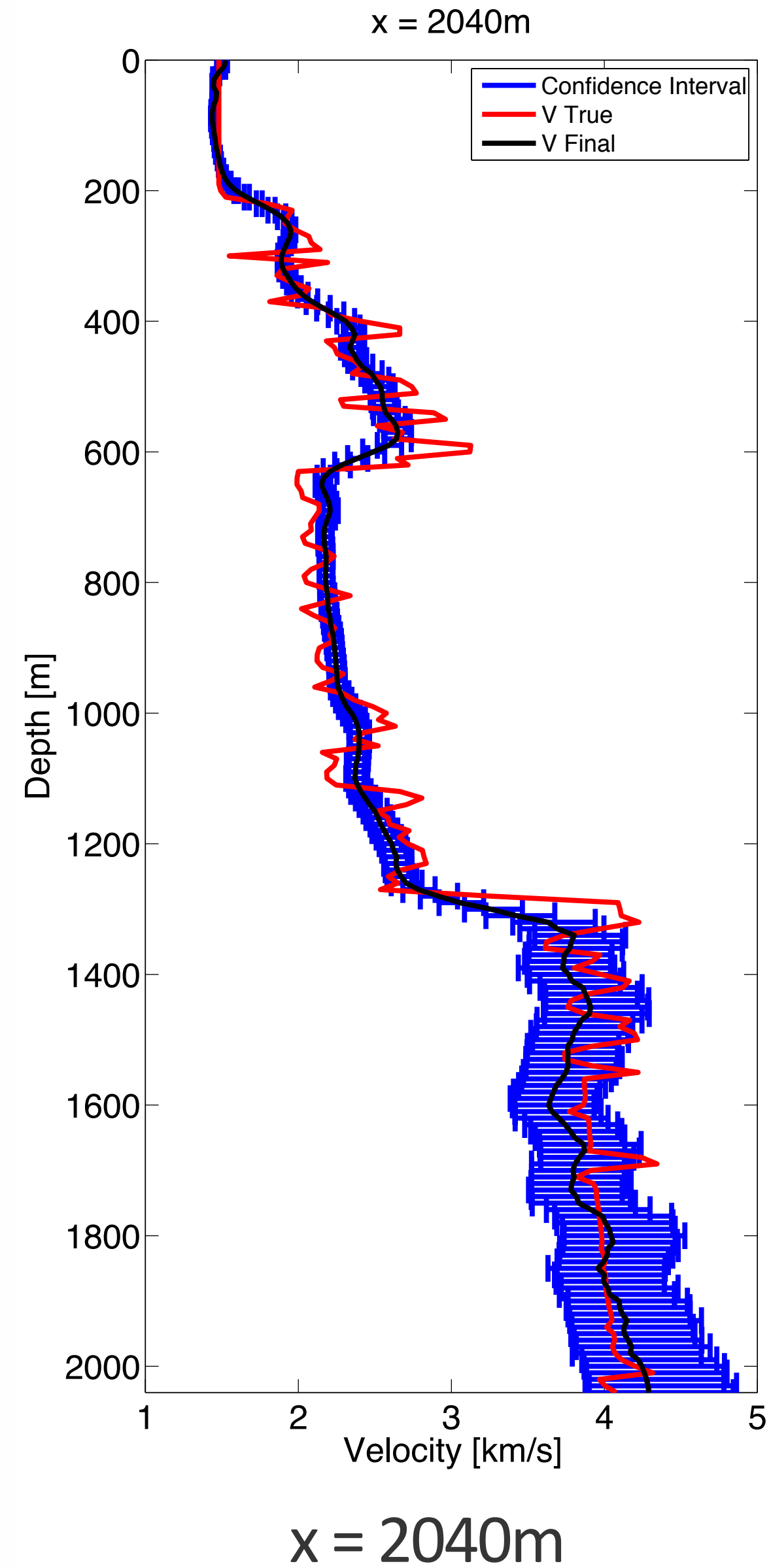
Which one is better?

Misfit?
Eyeball norm?

Uncertainty?
Standard Deviation?

Confidence interval

Fast Accurate!!



Bayesian theory

Deterministic inverse problem:

$$\mathbf{m}^* = \arg \min \left(\frac{1}{2N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{i\text{obs}}\|_{W_i}^2 + \frac{1}{2} \|\mathbf{m} - \overline{\mathbf{m}}\|_R^2 \right)$$

Statistical inverse problem with Bayesian theory:

$$\pi_{post}(\mathbf{m}) := \pi(\mathbf{m}|\mathbf{d}_{obs}) \propto \pi_{prior}(\mathbf{m})\pi(\mathbf{d}_{obs}|\mathbf{m})$$

where \mathbf{m} is the model parameter, and \mathbf{d}_{obs} is the observed data

Bayesian theory

Assume:

$$\text{noise} \sim \mathcal{N}(0, \Gamma_{noise})$$

$$\text{prior model distribution} \sim \mathcal{N}(\mathbf{m}_{prior}, \Gamma_{prior}).$$

Negative log-posterior of the posterior pdf:

$$V(\mathbf{m}) := -\log \pi_{post}(\mathbf{m}) := \frac{1}{2N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{inoise}^{-1}}^2 + \frac{1}{2} \|\mathbf{m} - \mathbf{m}_{prior}\|_{\Gamma_{prior}^{-1}}^2$$

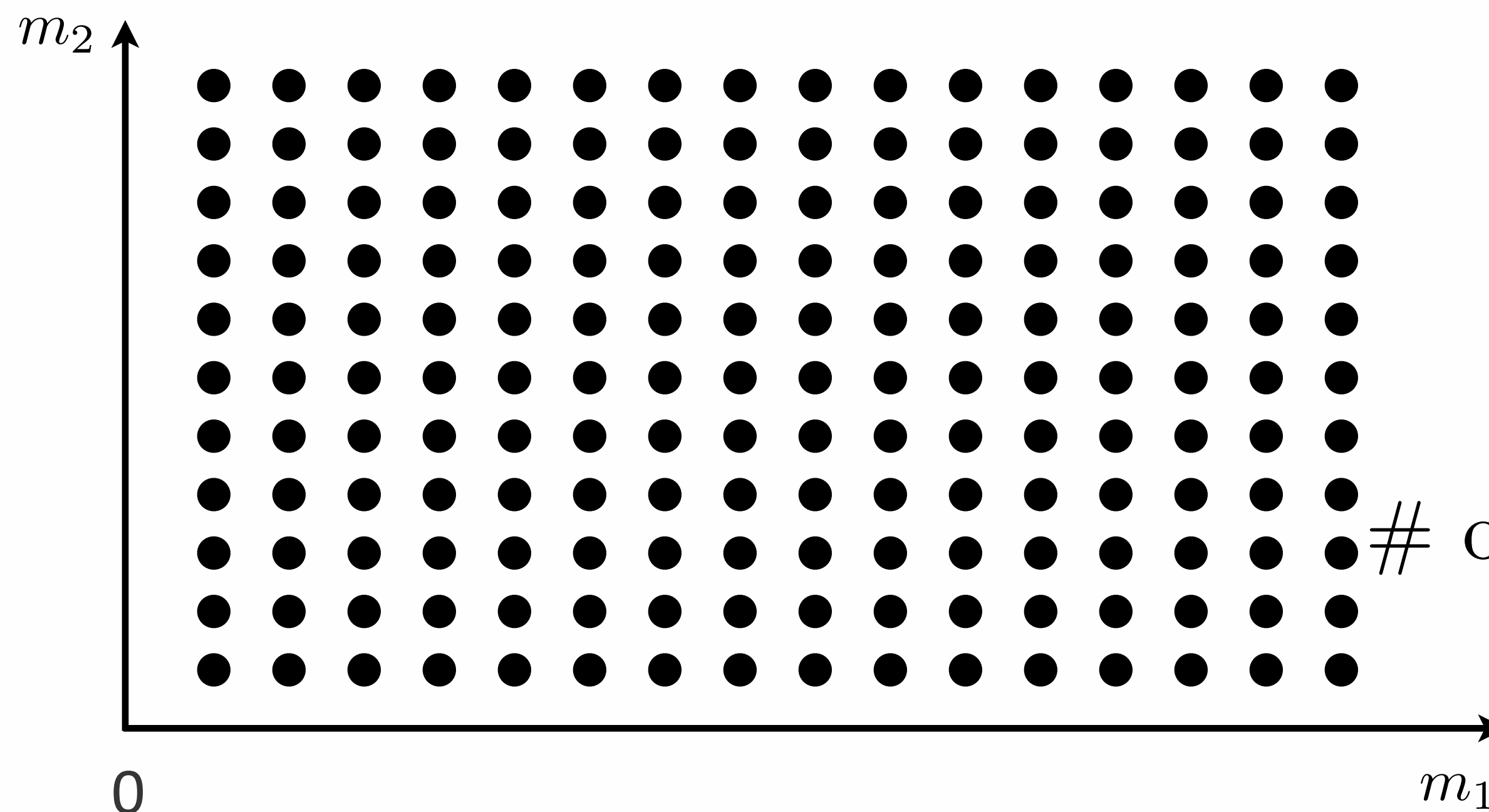
$$\mathbf{m}_{MAP} = \mathbf{m}^*$$

$$\|\mathbf{x}\|_{\Gamma_{noise}^{-1}}^2 := \mathbf{x}^T \Gamma_{noise}^{-1} \mathbf{x}$$

Sampling the posterior pdf

How to obtain the posterior probability density function?

- Compute $\pi_{post}(\mathbf{m})$ by discretization? $N = N_{sub_sample} * N_{para}$



$$N_{para} \gg 10^6$$
$$N_{sub_sample} \gg 10^2$$

$$N \gg 10^8$$

$$N_s \gg 10^2$$

$$\# \text{ of PDE solvers} \gg N * N_s$$

Approximate the pdf

How to obtain the posterior probability density function?

- Markov chain Monte Carlo method ? Metropolis - Hasting method

At sample \mathbf{m}_k

Draw sample \mathbf{y} from the proposal distribution $\tilde{\pi}_k(\mathbf{m})$

if $\min(1, \frac{\pi_{post}(\mathbf{y})\tilde{\pi}_y(\mathbf{m}_k)}{\pi_{post}(\mathbf{m}_k)\tilde{\pi}_k(\mathbf{y})}) > \alpha$

 set $\mathbf{m}_{k+1} = \mathbf{y}$

else

 regenerate \mathbf{y}

end

Approximate the pdf

How to obtain the posterior probability density function?

- Stochastic Newton MCMC ?

$$V(\mathbf{m}) := -\log \pi_{post}(\mathbf{m}) := \frac{1}{2N_s} \sum_{i=1}^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{i\text{obs}}\|_{\Gamma_{i\text{noise}}^{-1}}^2 + \frac{1}{2} \|\mathbf{m} - \mathbf{m}_{prior}\|_{\Gamma_{prior}^{-1}}^2$$

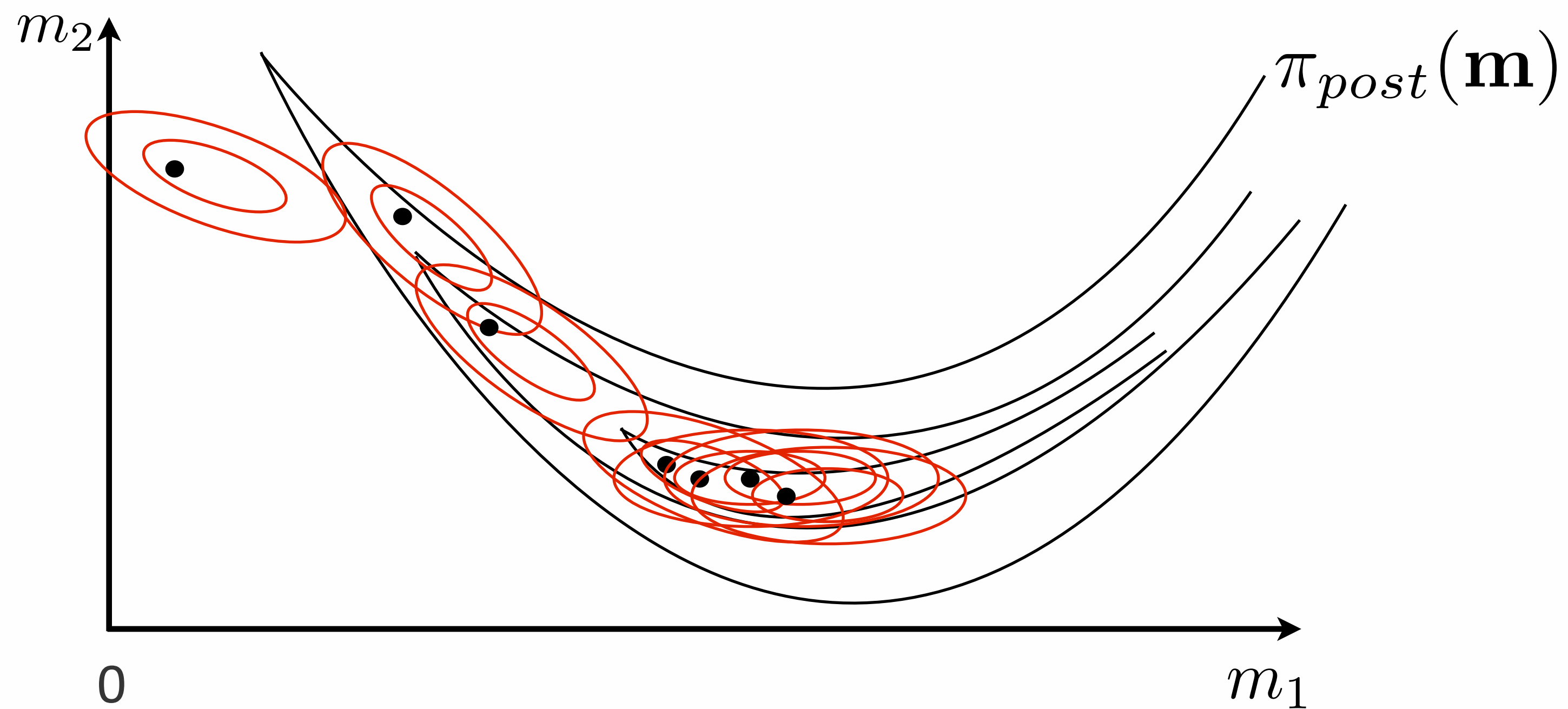
$$\tilde{\pi}_k(\mathbf{m}) \sim \mathcal{N}(\mathbf{m}_k - \mathbf{H}_k^{-1} \mathbf{g}_k, \mathbf{H}_k^{-1})$$

$$\begin{bmatrix} \mathbf{L}^T \end{bmatrix}_{n \times n} \begin{bmatrix} \mathbf{H}_{\text{misfit}} \end{bmatrix}_{n \times n} \begin{bmatrix} \mathbf{L} \end{bmatrix}_{n \times n} \approx \begin{bmatrix} \mathbf{V}_r \end{bmatrix}_{n \times r} \begin{bmatrix} \mathbf{D}_r \end{bmatrix}_{r \times r} \begin{bmatrix} \mathbf{V}_r^T \end{bmatrix}_{r \times n}$$

$r \ll n$
(James Martin *et al*, 2012)

Approximate the pdf

$$\tilde{\pi}_k(\mathbf{m}) \sim \mathcal{N}(\mathbf{m}_k - \mathbf{H}_k^{-1} \mathbf{g}_k, \mathbf{H}_k^{-1})$$



Approximate the pdf

Challenges:

- Low - rank approximation may not be correct.
- The computational cost of estimating Hessian is huge.
- The computational cost of calculating the posterior distribution probability density function is huge.

$$\text{Computational Cost} \sim \mathcal{O}(N_{\text{sample}} * N_s)$$

BFGS Hessian

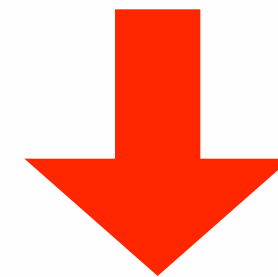
At the stage k : approximated Hessian \mathbf{B}_k and gradient \mathbf{g}_k ,
Update \mathbf{B}_k by adding two rank one matrices:

$$\mathbf{B}_{k+1} = \mathbf{B}_k + \mathbf{U}_k + \mathbf{V}_k$$

where \mathbf{B}_{k+1} should satisfy

$$\mathbf{B}_{k+1}(\mathbf{x}_{k+1} - \mathbf{x}_k) = \mathbf{g}_{k+1} - \mathbf{g}_k$$

No additional computational cost!!!



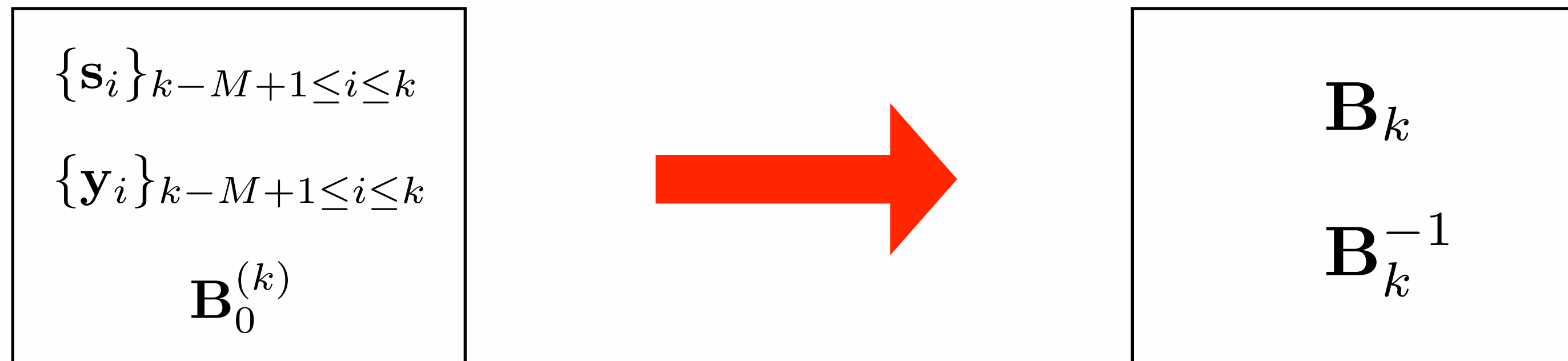
$$\begin{aligned} \mathbf{y}_k &= \mathbf{g}_{k+1} - \mathbf{g}_k \\ \mathbf{s}_k &= \mathbf{x}_{k+1} - \mathbf{x}_k, \end{aligned}$$

$$\mathbf{B}_{k+1} = \mathbf{B}_k + \frac{\mathbf{y}_k \mathbf{y}_k^T}{\mathbf{y}_k^T \mathbf{s}_k} - \frac{\mathbf{B}_k \mathbf{s}_k \mathbf{s}_k^T \mathbf{B}_k}{\mathbf{s}_k^T \mathbf{B}_k \mathbf{s}_k}$$

$$\mathbf{B}_{k+1}^{-1} = \left(\mathbf{I} - \frac{\mathbf{s}_k \mathbf{y}_k^T}{\mathbf{y}_k^T \mathbf{s}_k} \right) \mathbf{B}_k^{-1} \left(\mathbf{I} - \frac{\mathbf{y}_k \mathbf{s}_k^T}{\mathbf{y}_k^T \mathbf{s}_k} \right) + \frac{\mathbf{s}_k \mathbf{s}_k^T}{\mathbf{y}_k^T \mathbf{s}_k}$$

I-BFGS Hessian with pseudo GN Hessian as a starter

Limited memory BFGS Hessian:



Choice of $\mathbf{B}_0^{(k)}$: pseudo GN Hessian

Randomized source subsampling

To speed up the computation of posterior pdf:

$$V(\mathbf{m}) := -\log \pi_{post}(\mathbf{m}) := \frac{1}{2N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 + \frac{1}{2} \|\mathbf{m} - \mathbf{m}_{prior}\|_{\Gamma_{prior}^{-1}}^2$$

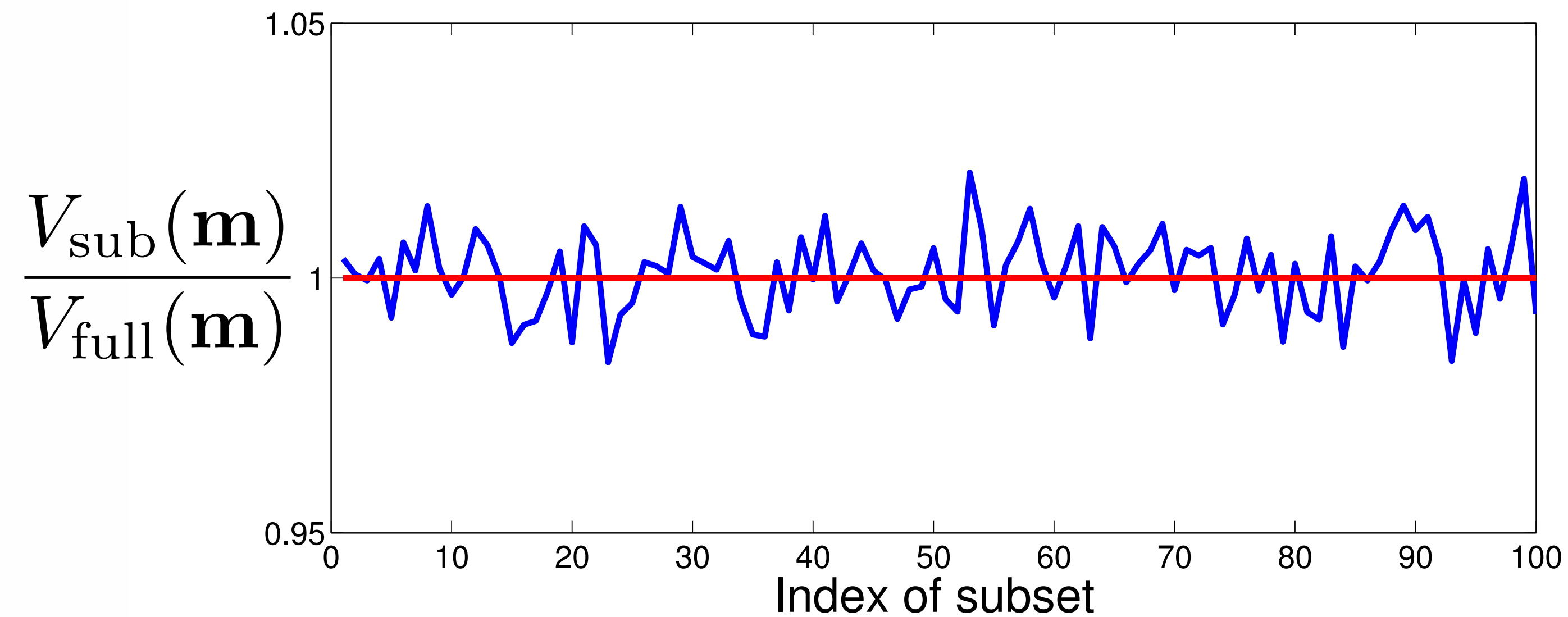
Randomized source subsampling:

$$\frac{1}{N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 = \frac{1}{\|\mathcal{I}_s\|} \sum_{i \in \mathcal{I}_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 + \epsilon$$

Randomized source subsampling

Computational Cost: $\mathcal{O}(N_s) \rightarrow \mathcal{O}(N_{rs}), N_{rs} \ll N_s$

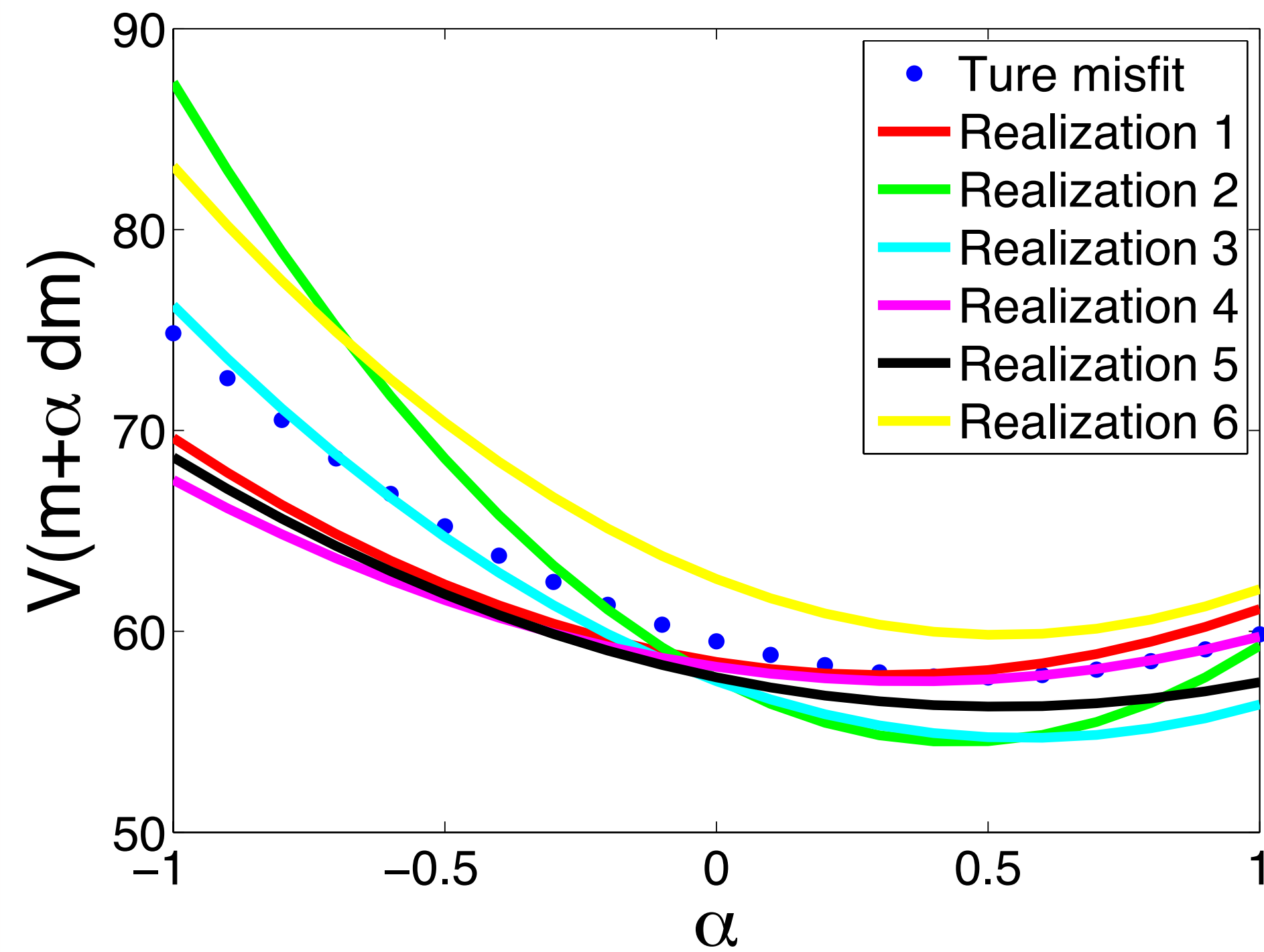
Subset misfit vs Full misfit



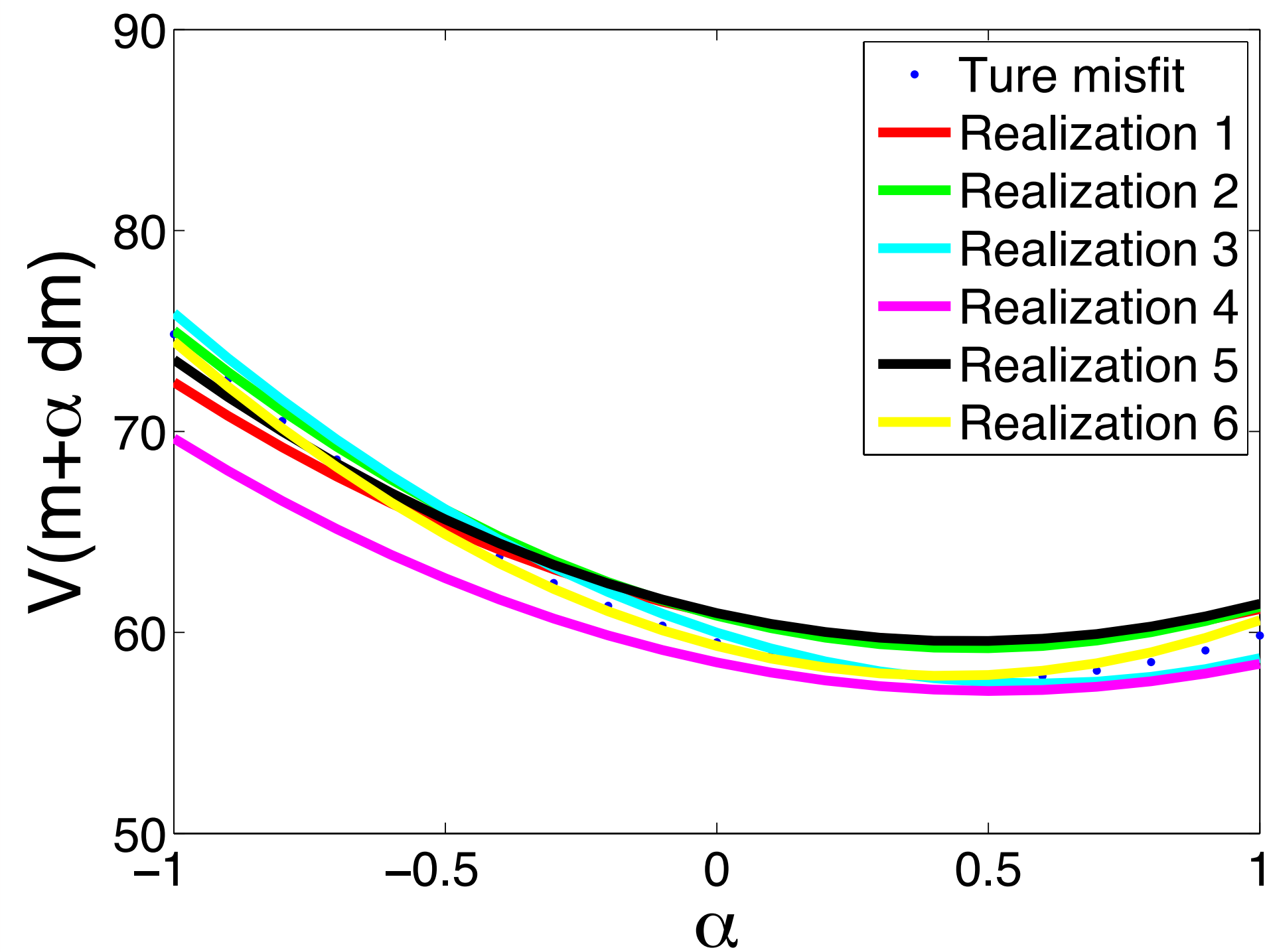
Randomized source subsampling

Computational Cost: $\mathcal{O}(N_s) \rightarrow \mathcal{O}(N_{rs})$

5 / 91 shots

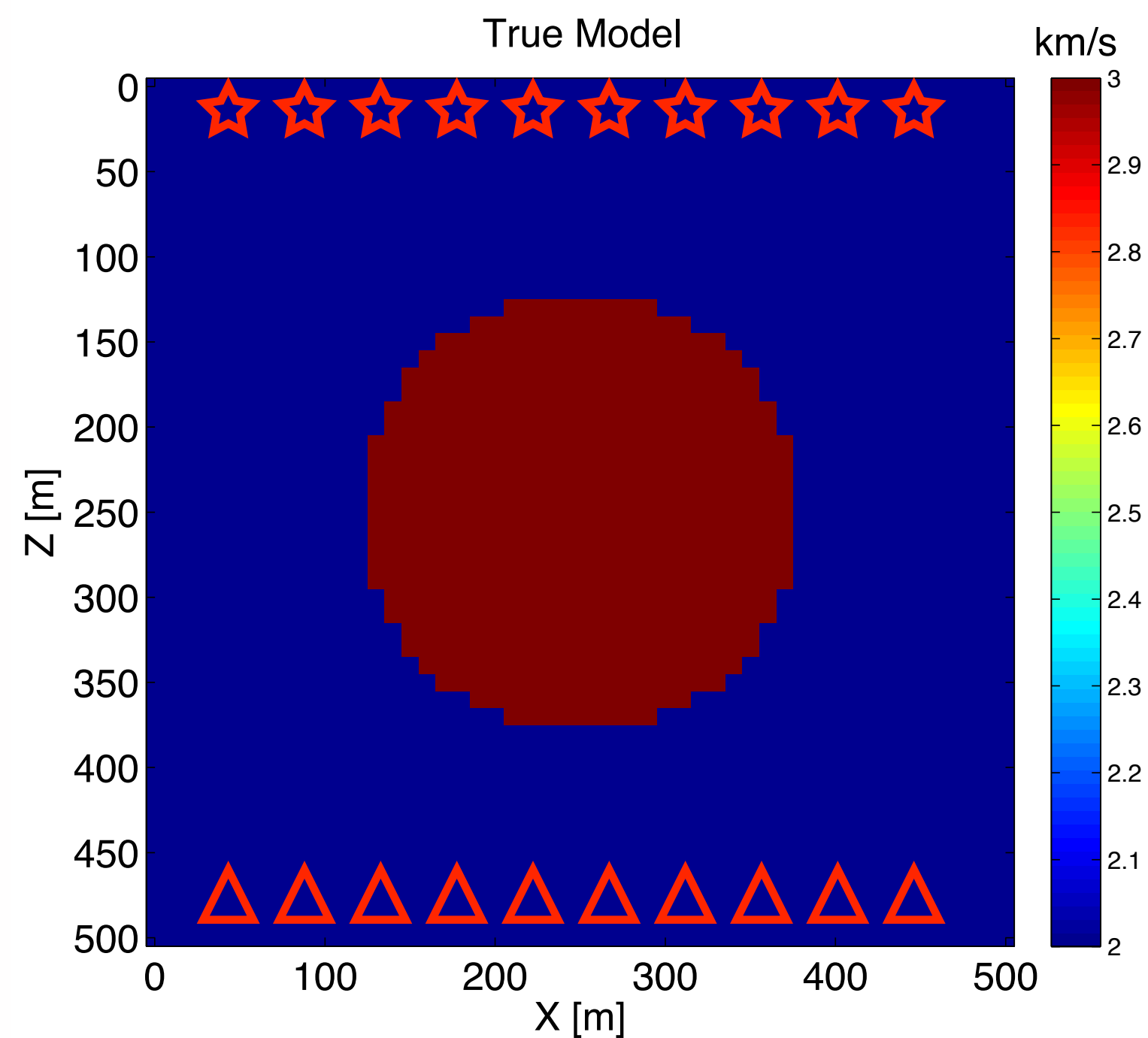


20 / 91 shots



Numerical Experiments

Camambert model



☆ - source
△ - receiver

Statistical parameter to be inverted:

Standard deviation : σ

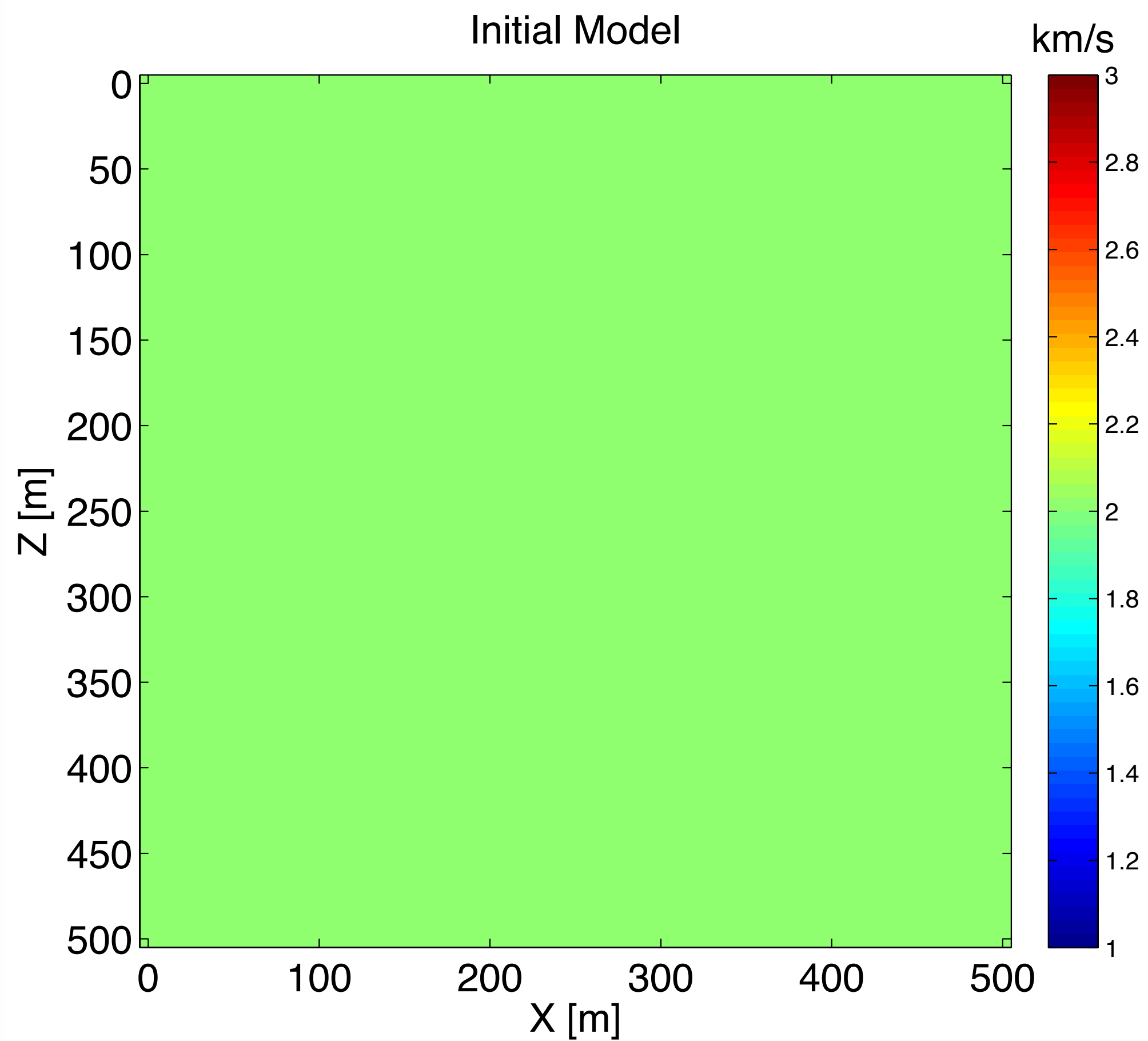
Confidence interval: $P(\mathbf{m} \in I_{ci}) \geq \alpha$

Acquisition Geometry:

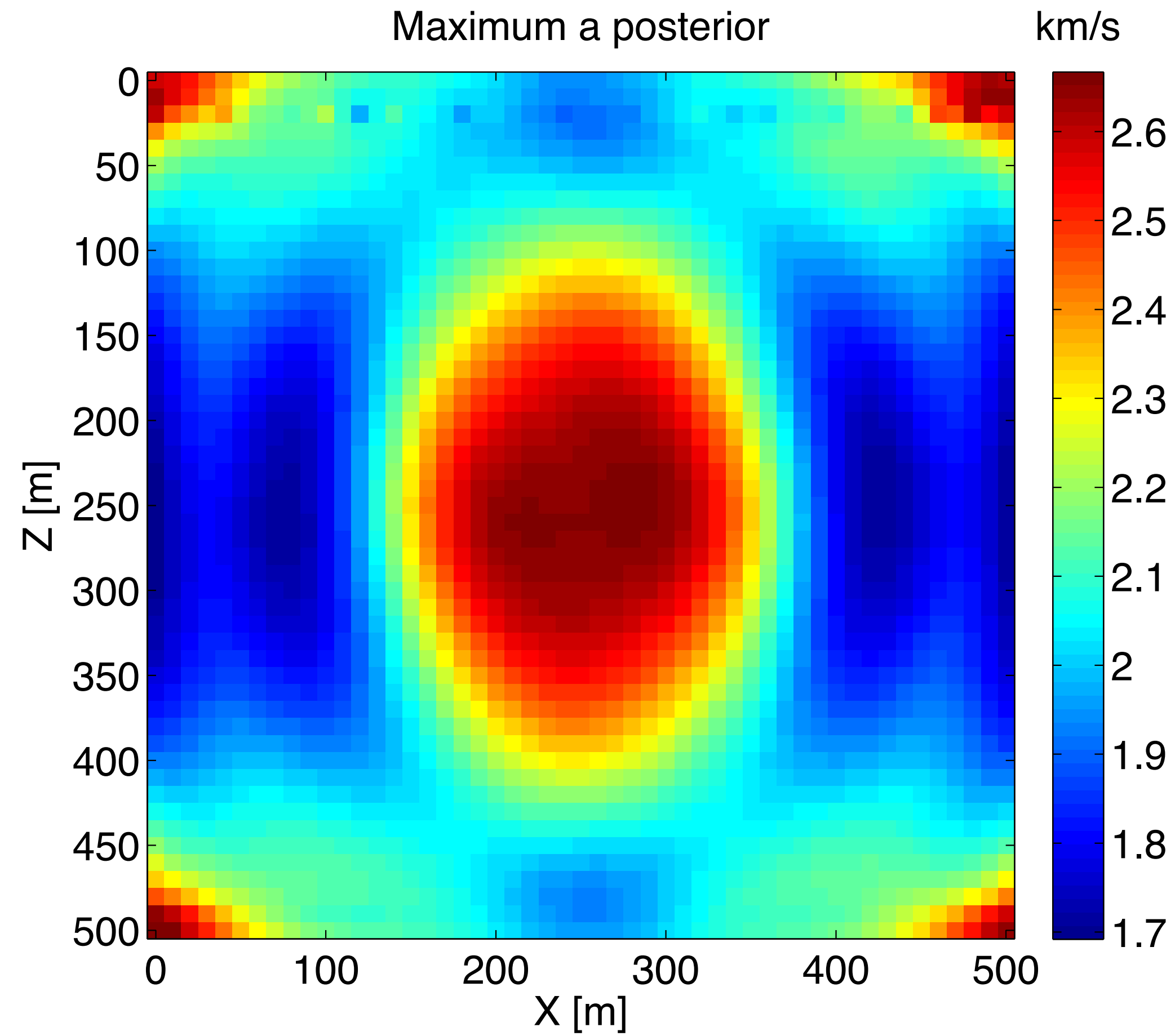
26 shots

51 receivers

10 frequencies

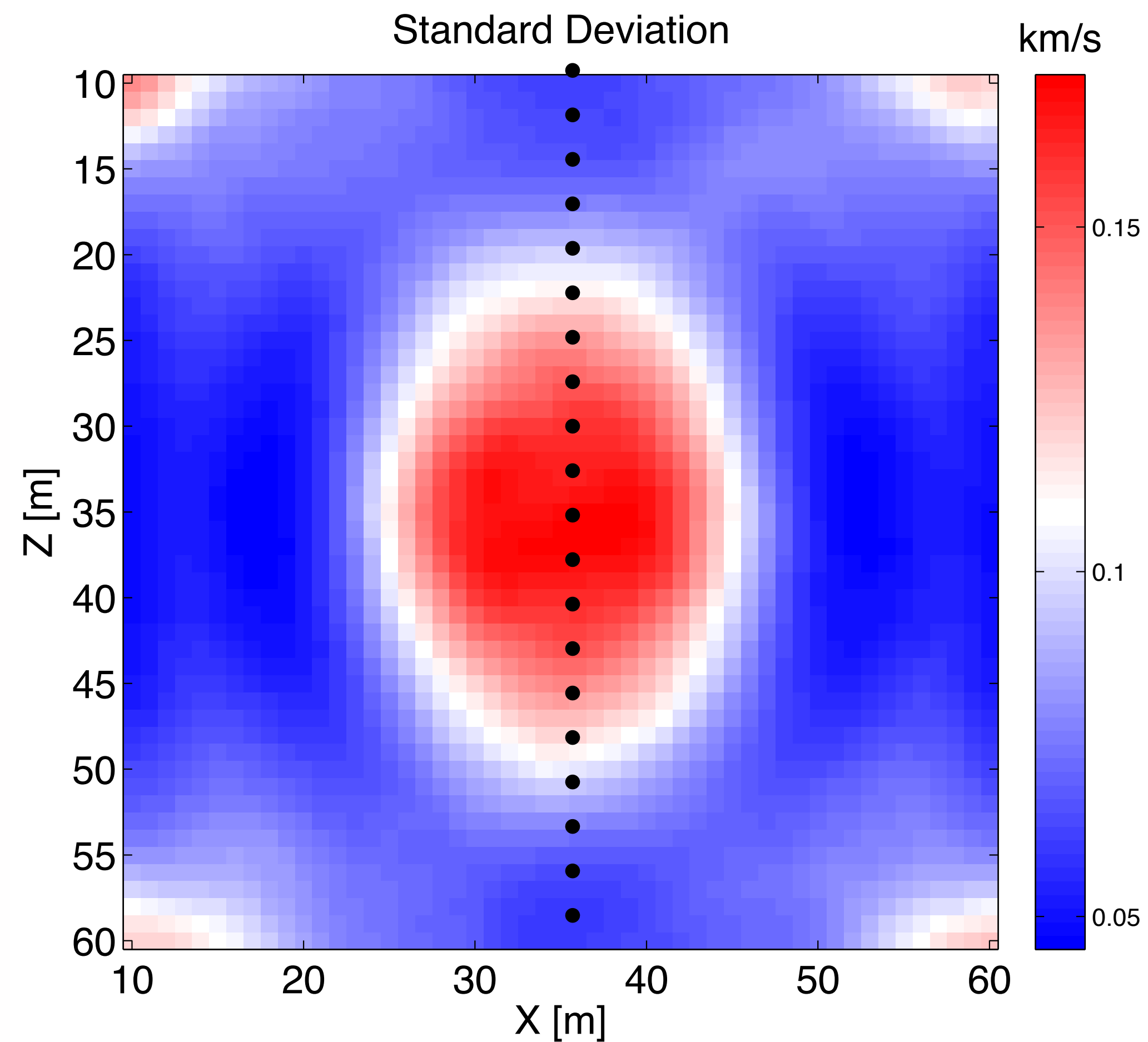


Initial model

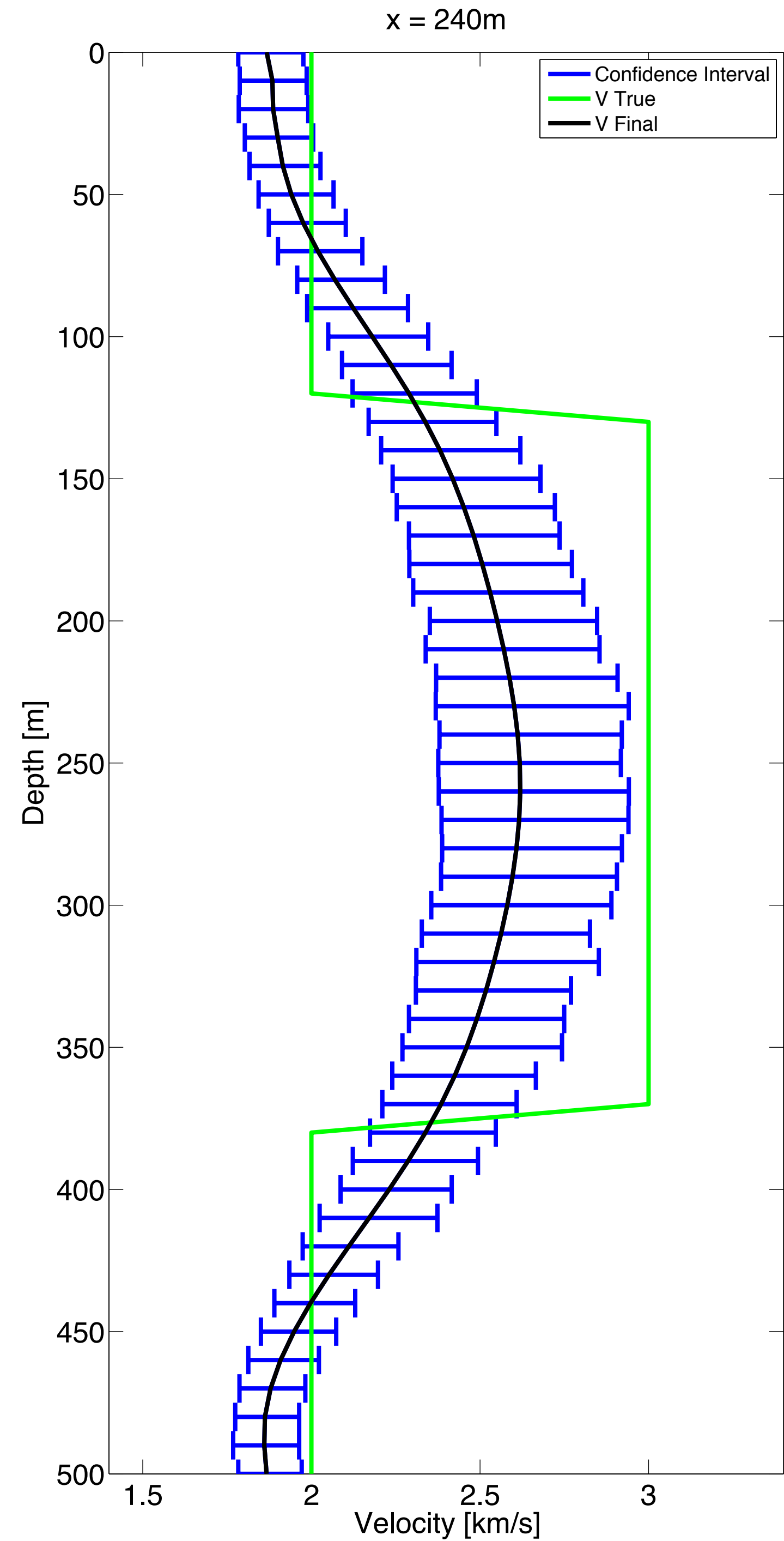


MAP

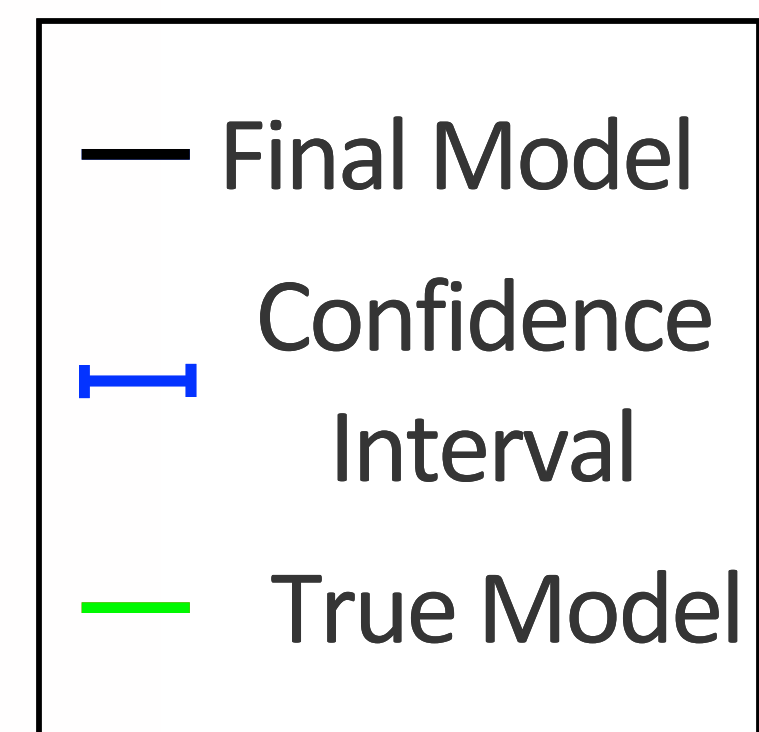
Nrs = 5



Standard Deviation

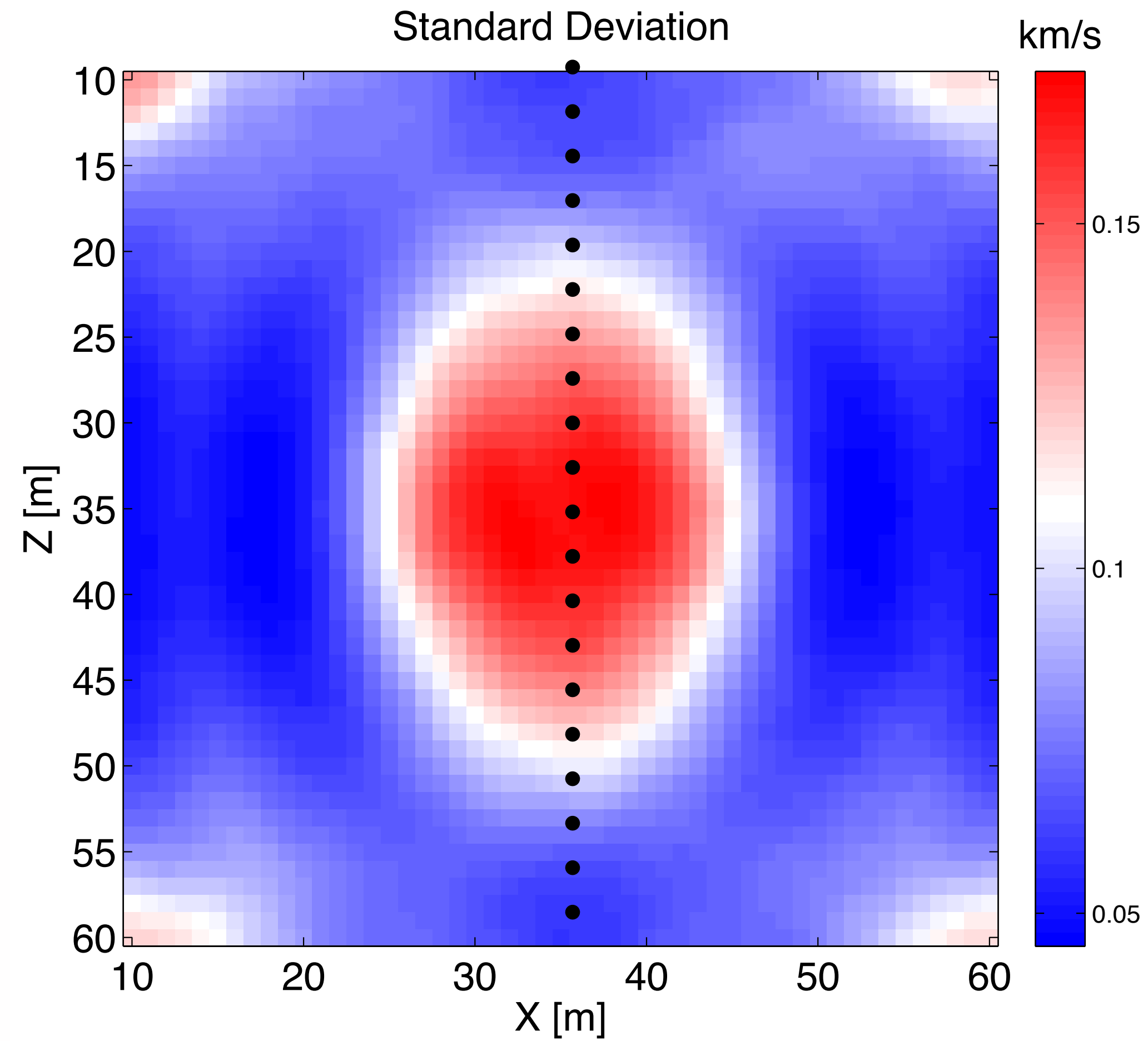


$$\alpha = 0.9$$

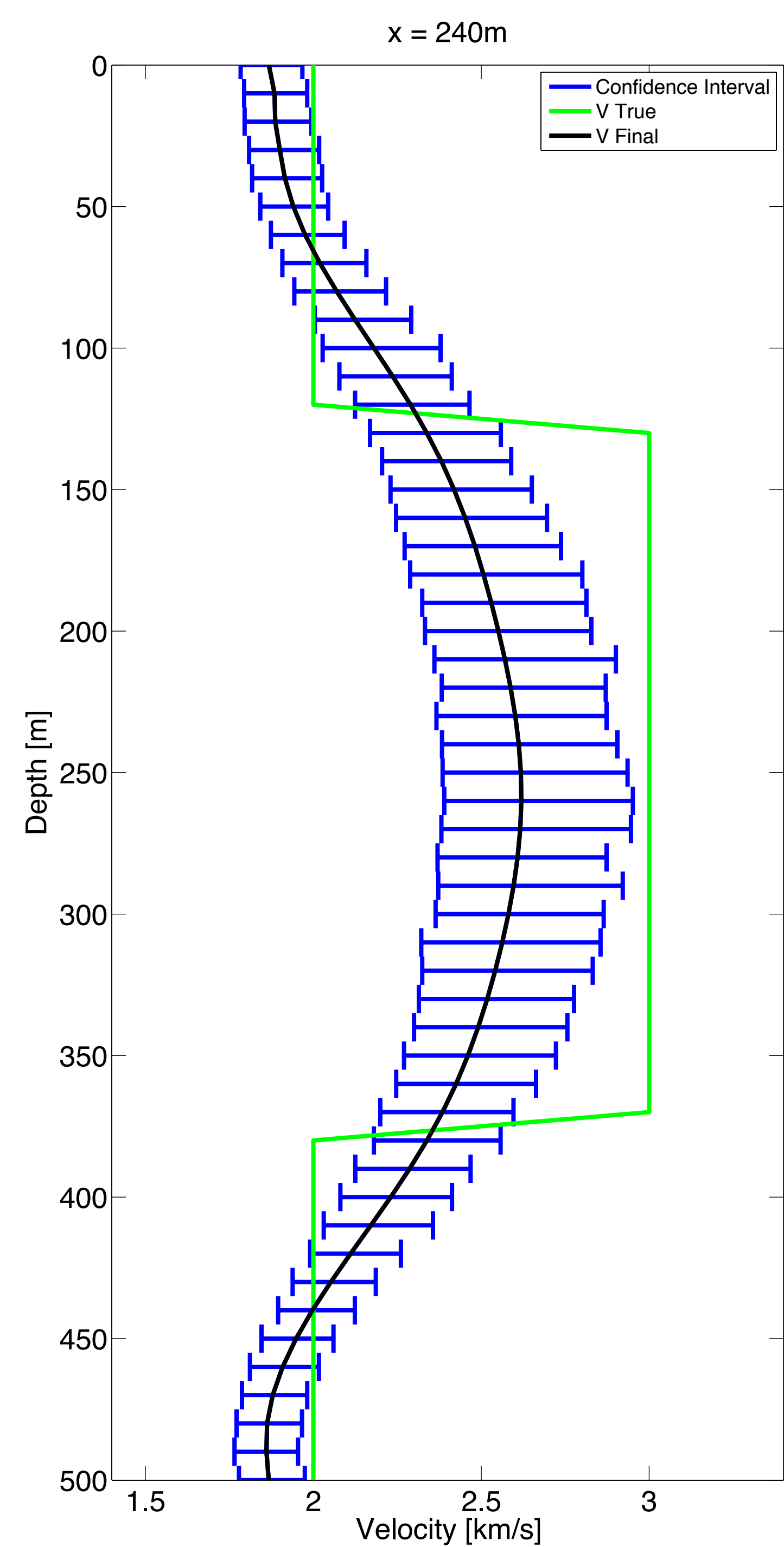


Confidence Interval

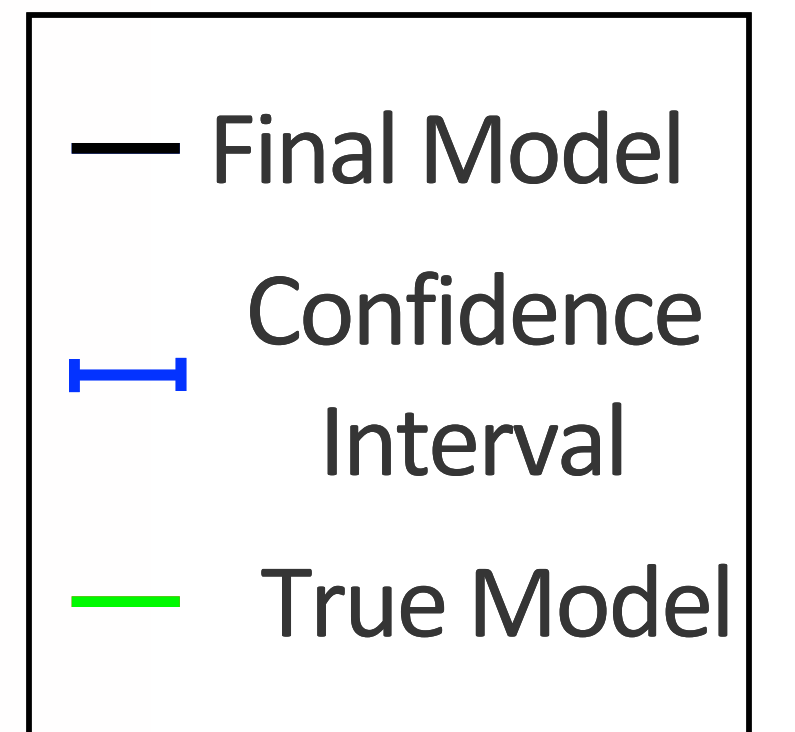
Nrs = 10



Standard Deviation

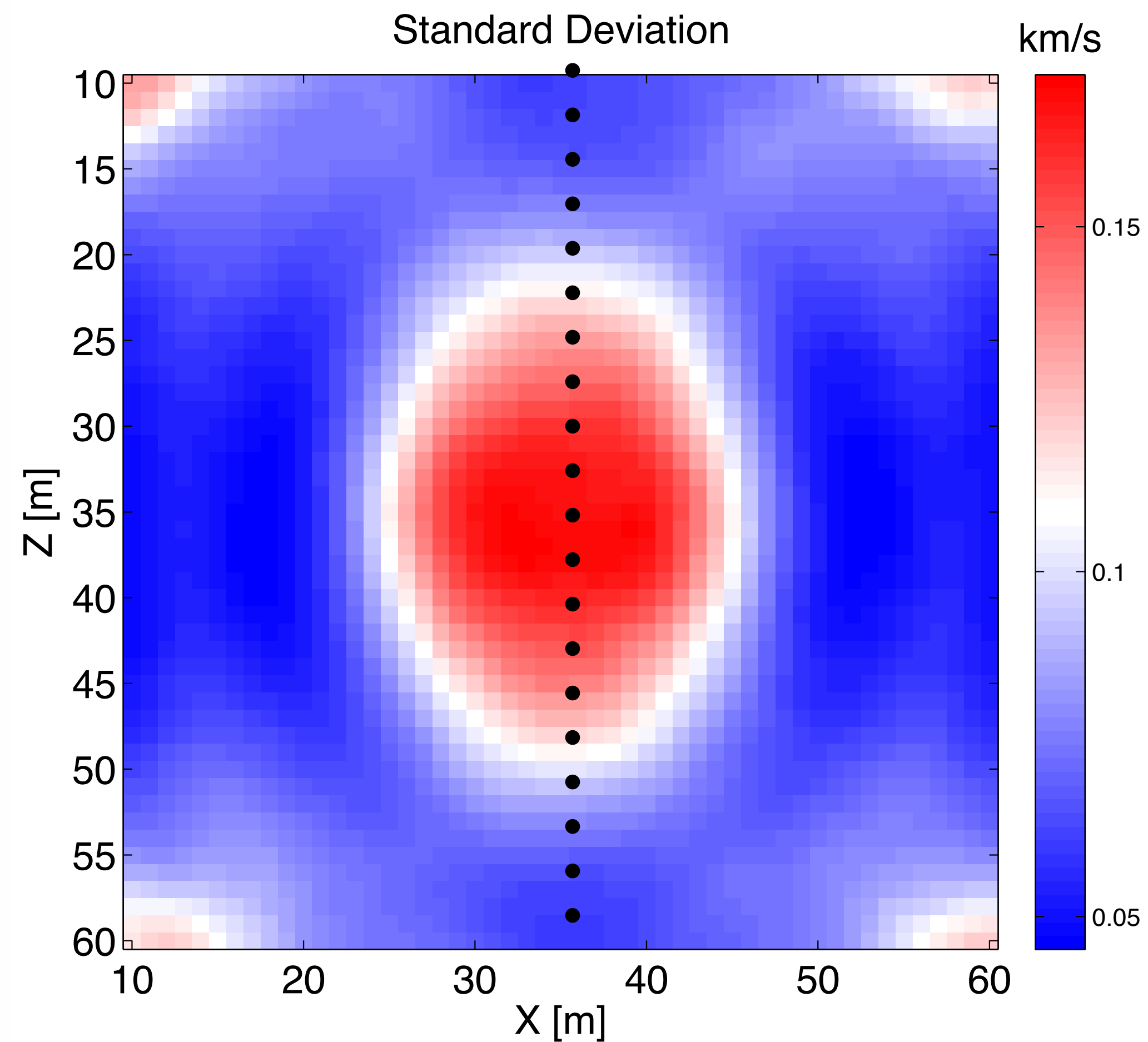


$$\alpha = 0.9$$

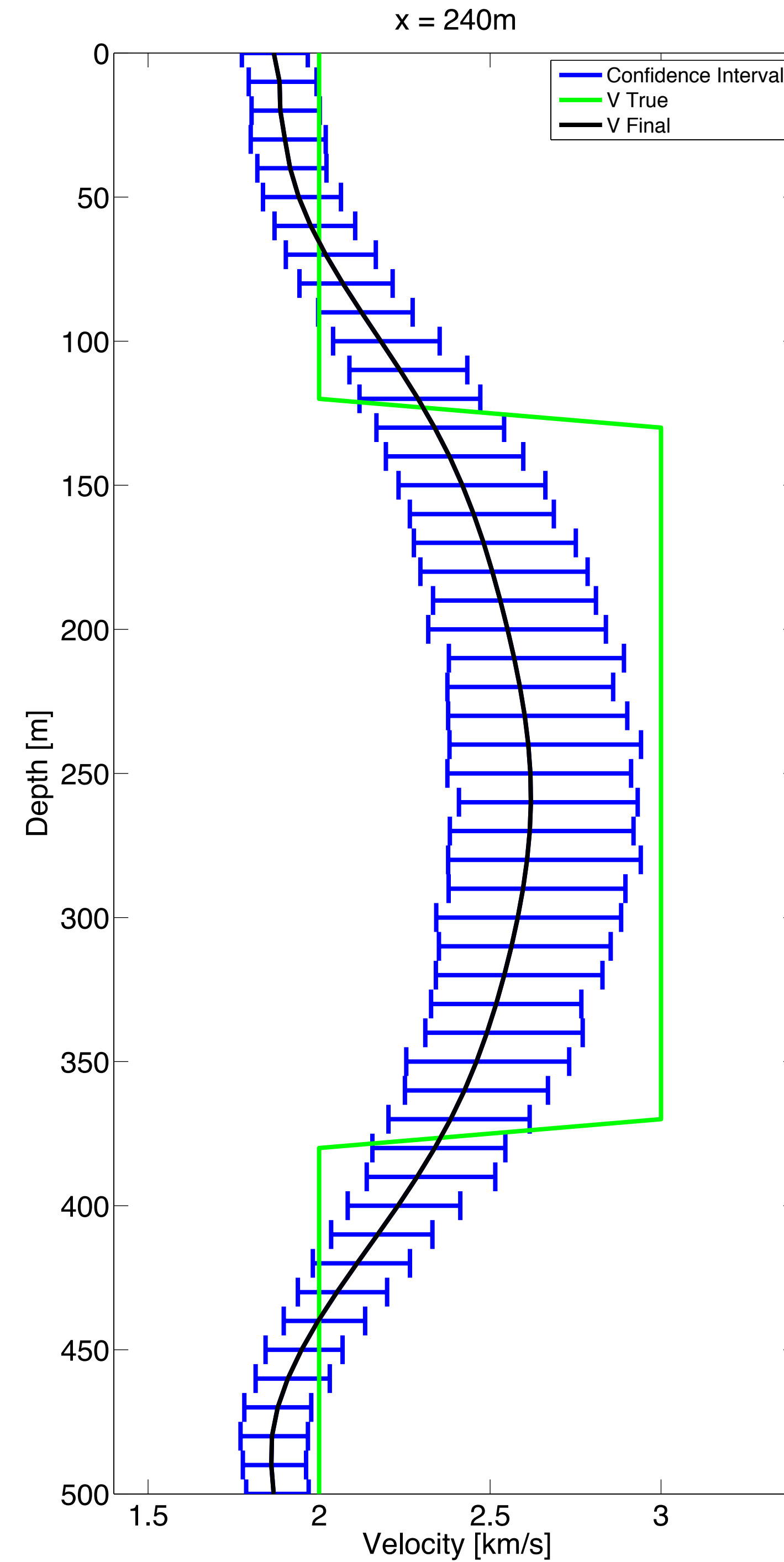


Confidence Interval

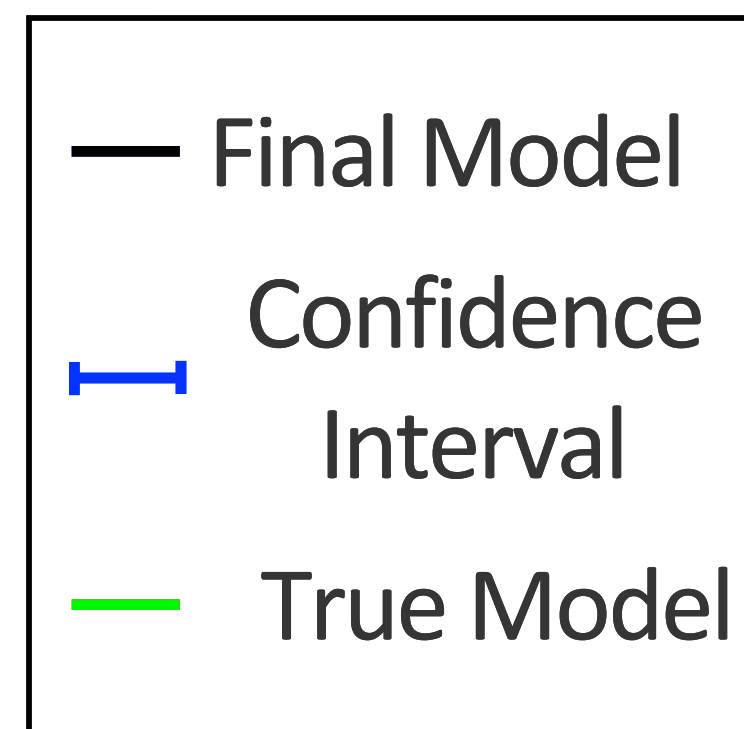
Nrs = 26



Standard Deviation



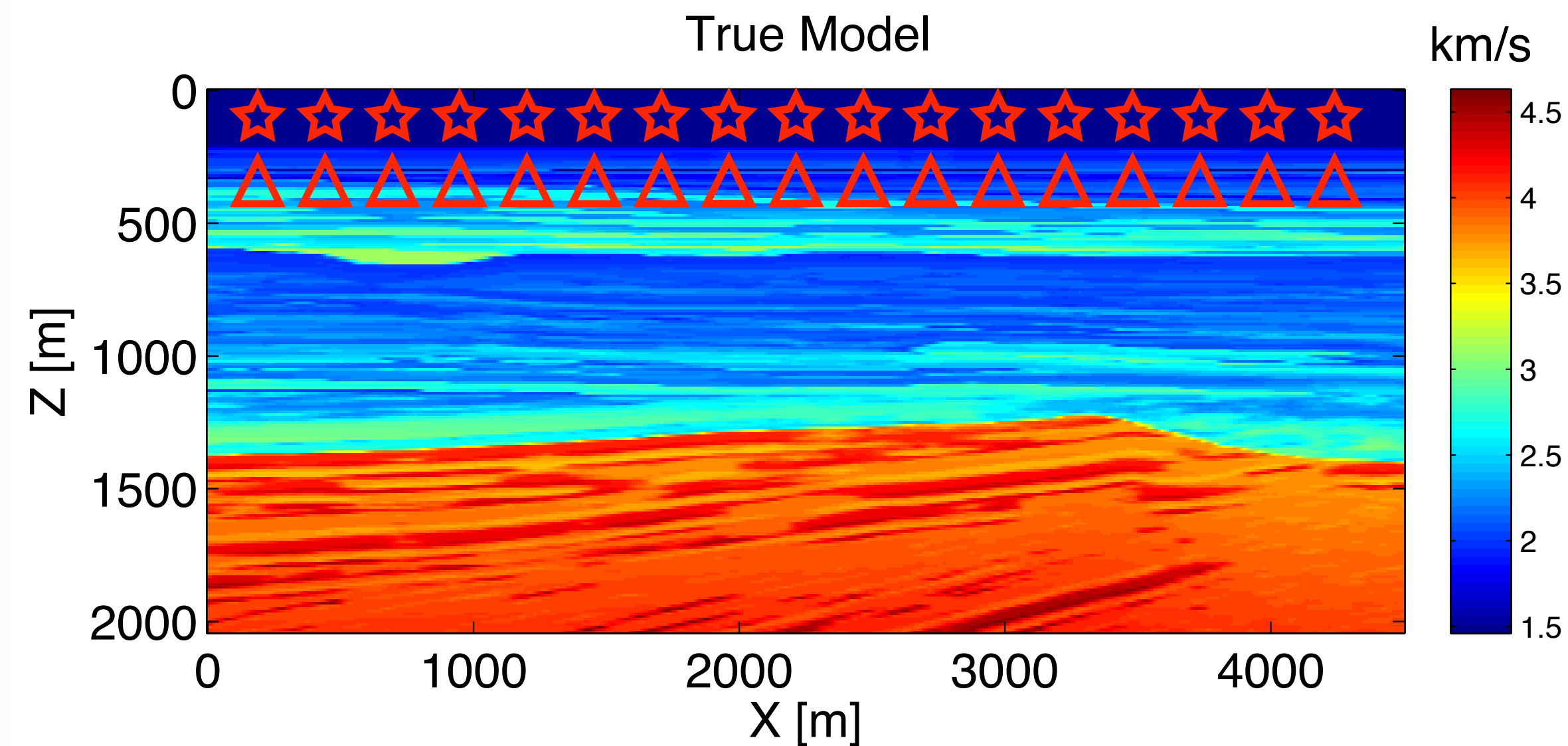
$$\alpha = 0.9$$



Confidence Interval

Numerical Experiments

BG model



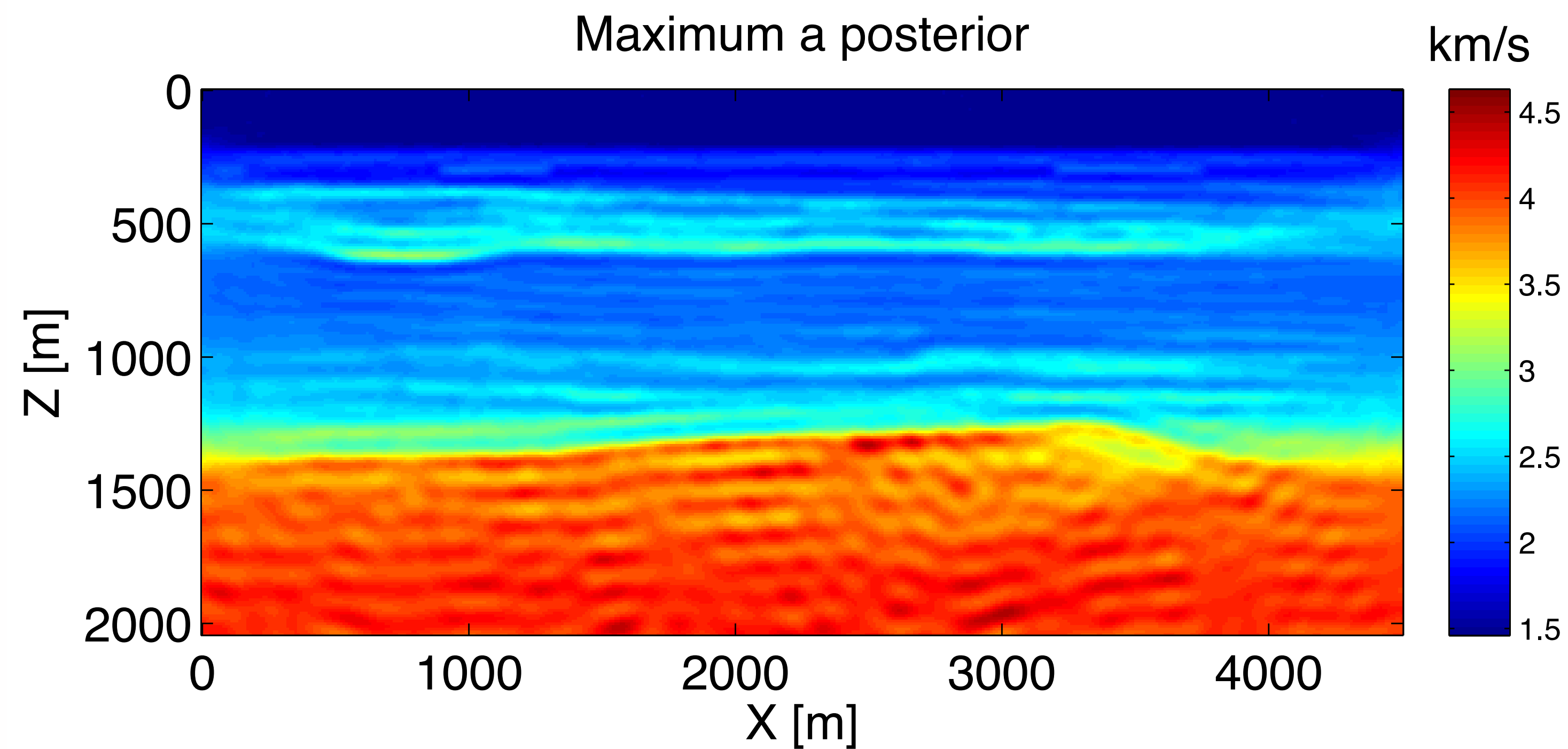
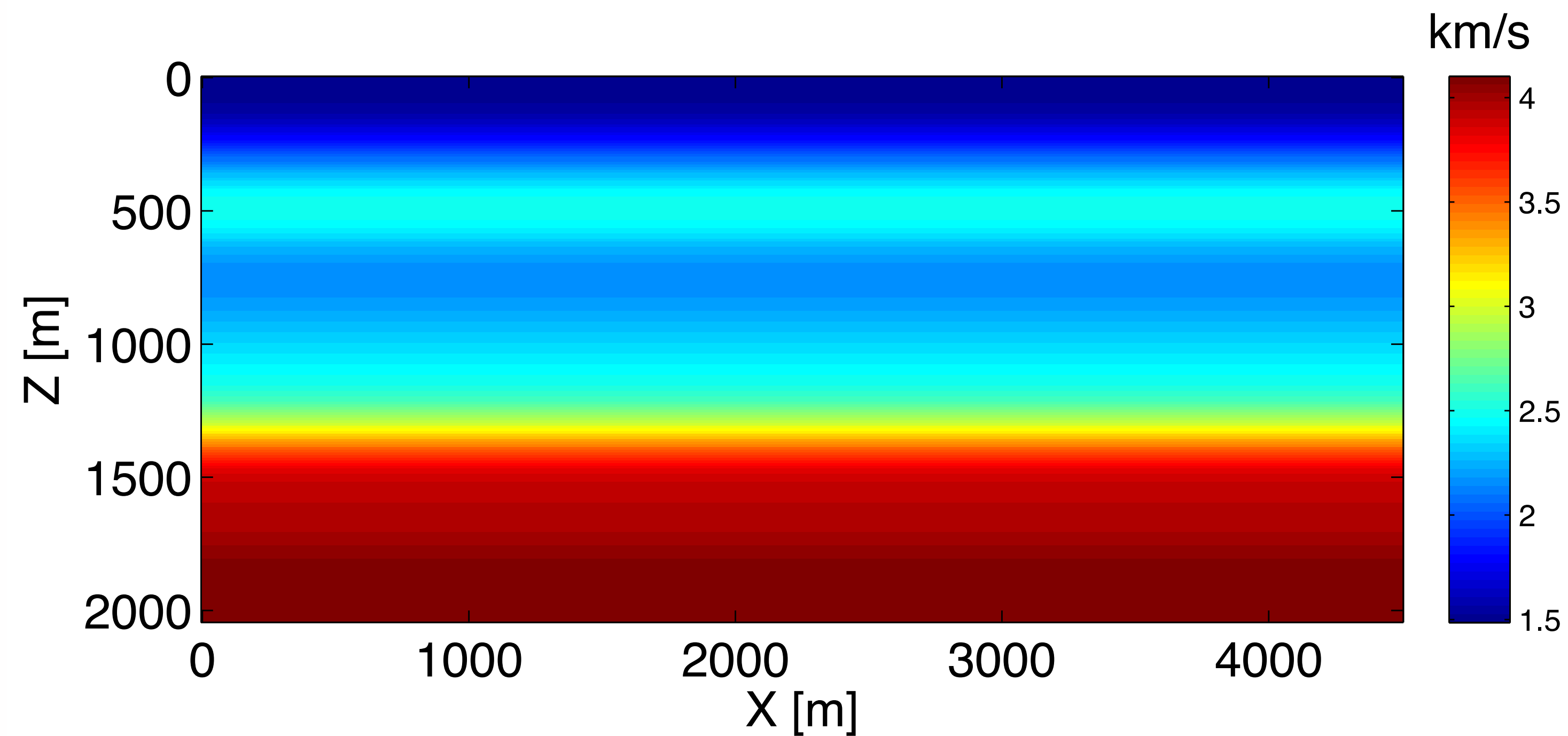
☆ - source
△ - receiver

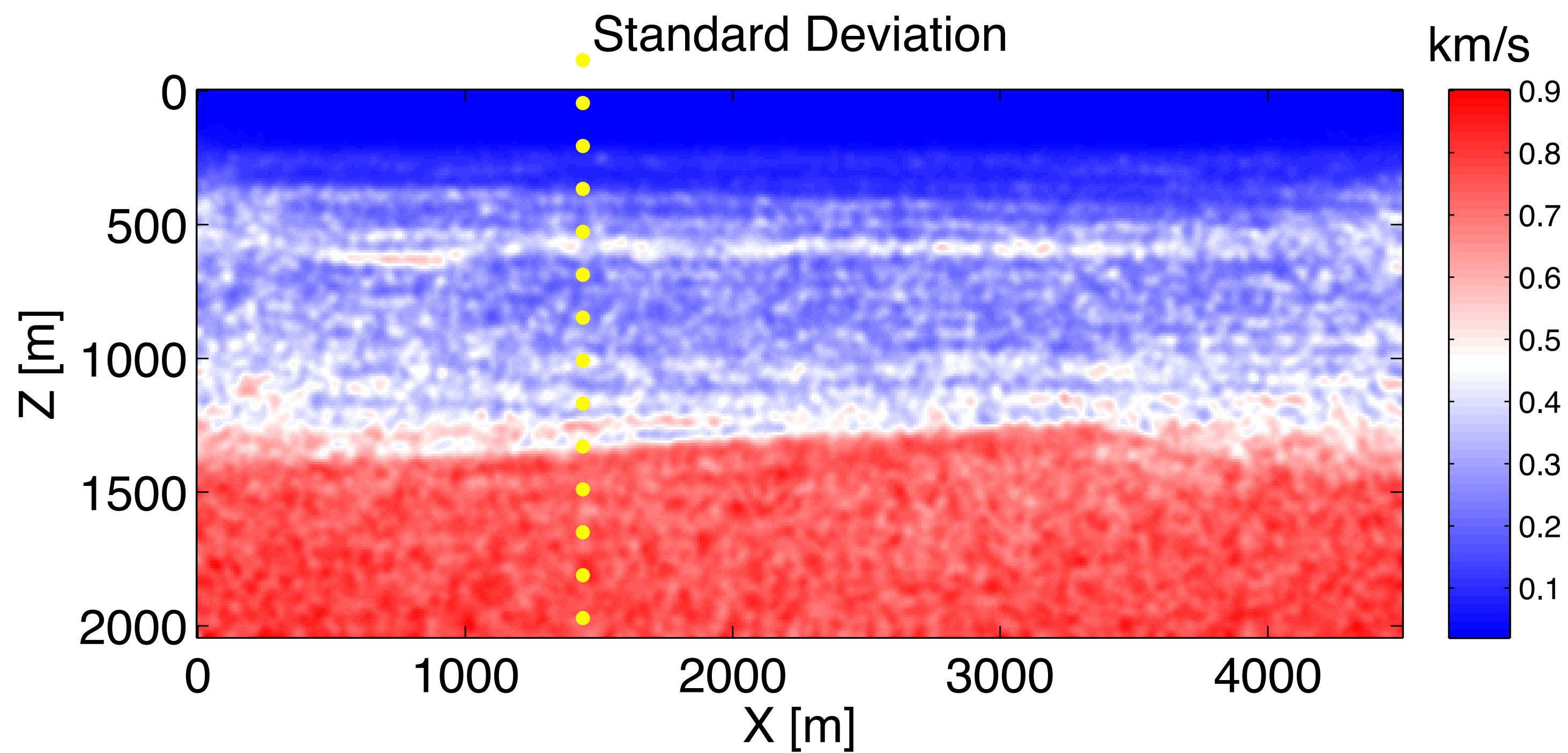
Acquisition Geometry:

91 shots

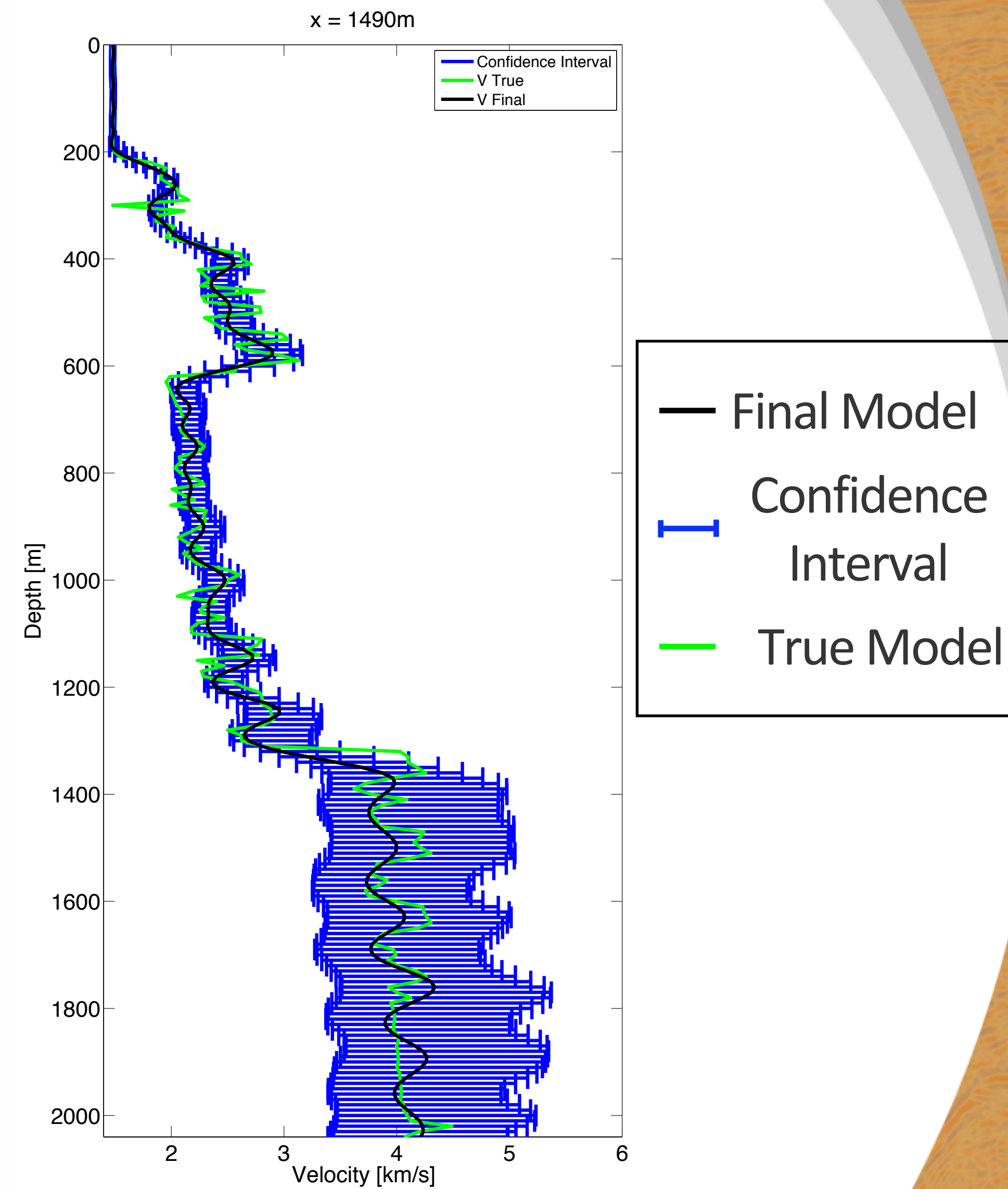
451 receivers

15 frequencies from 3Hz to 17Hz

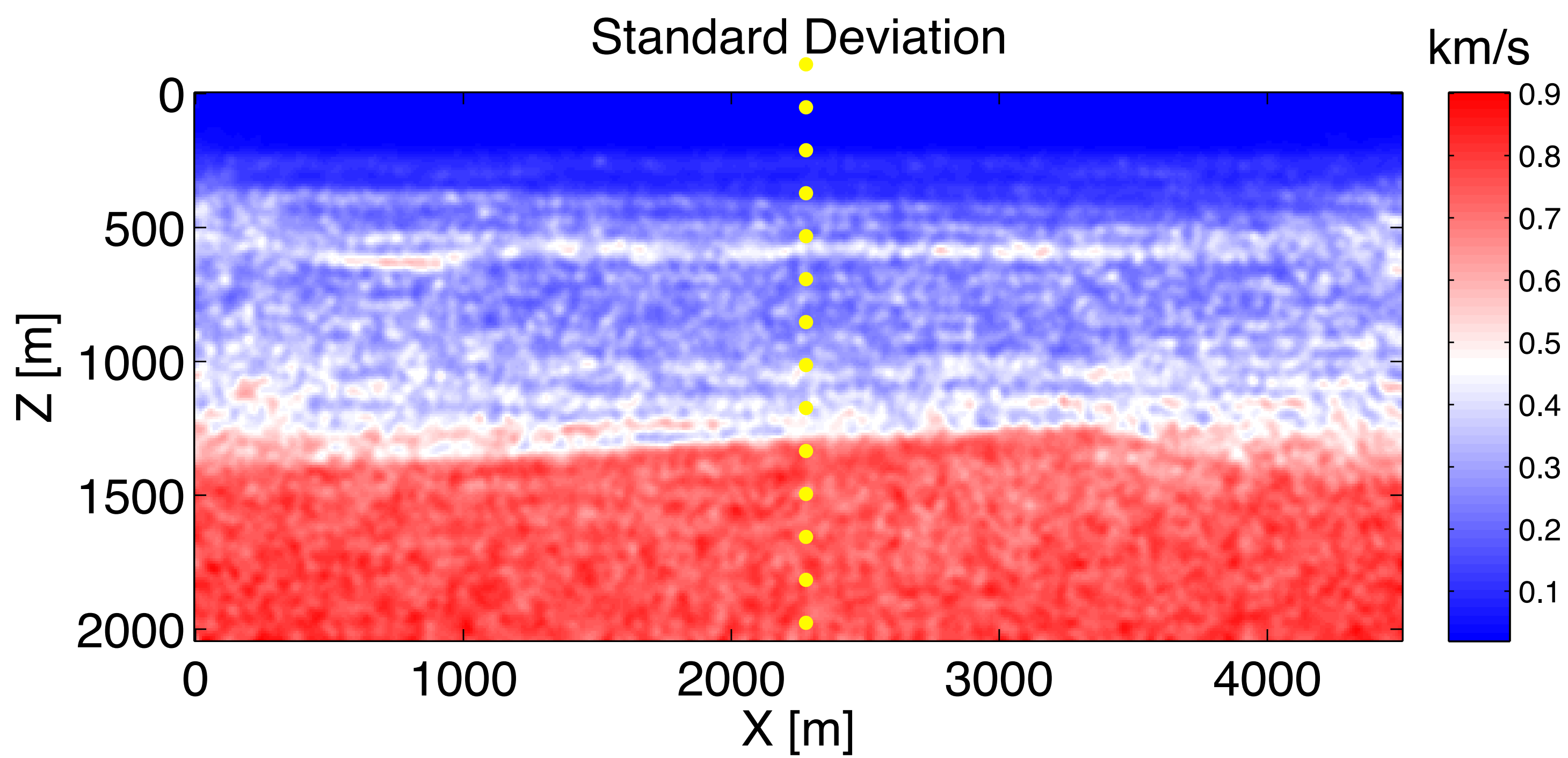




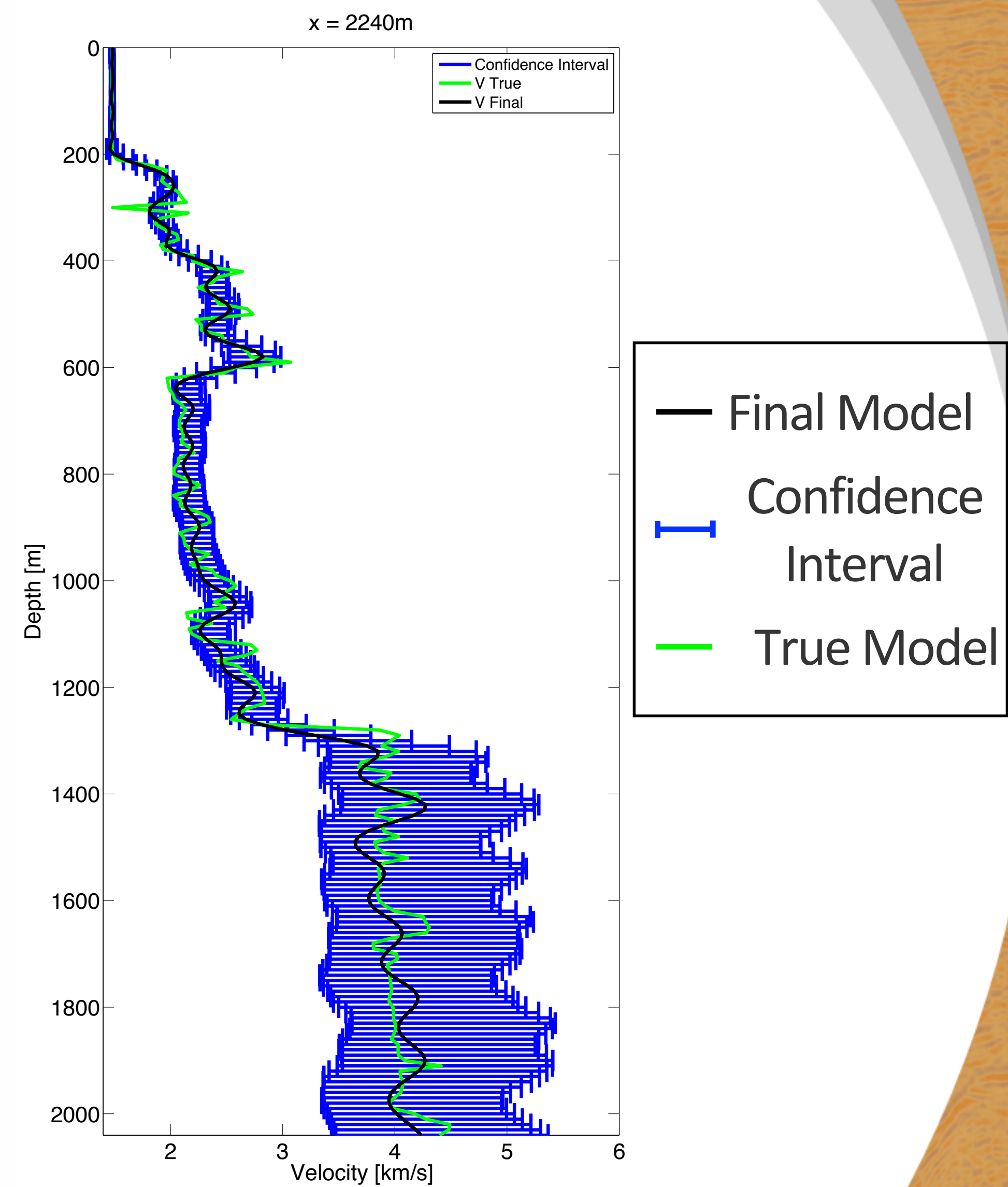
Standard deviation



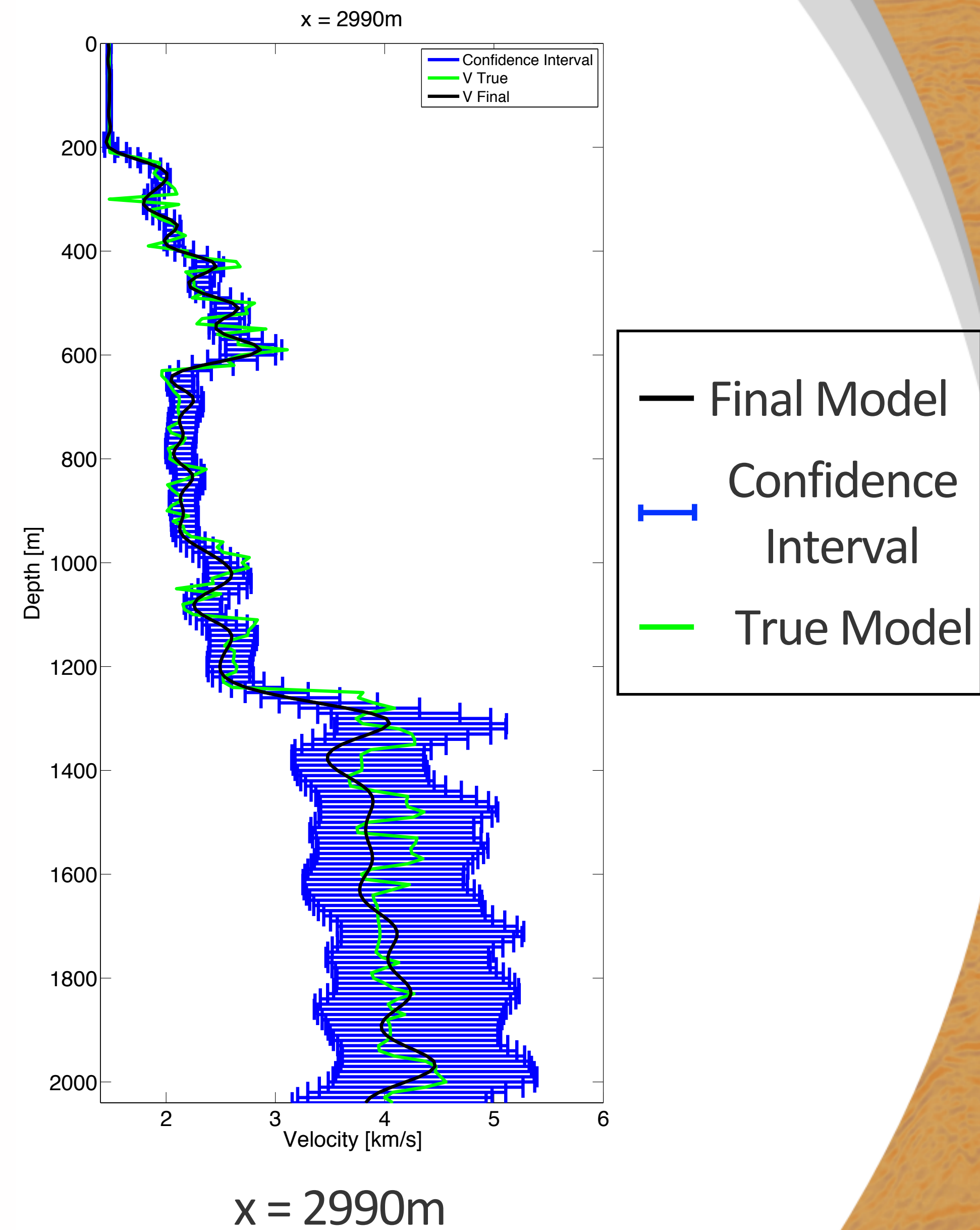
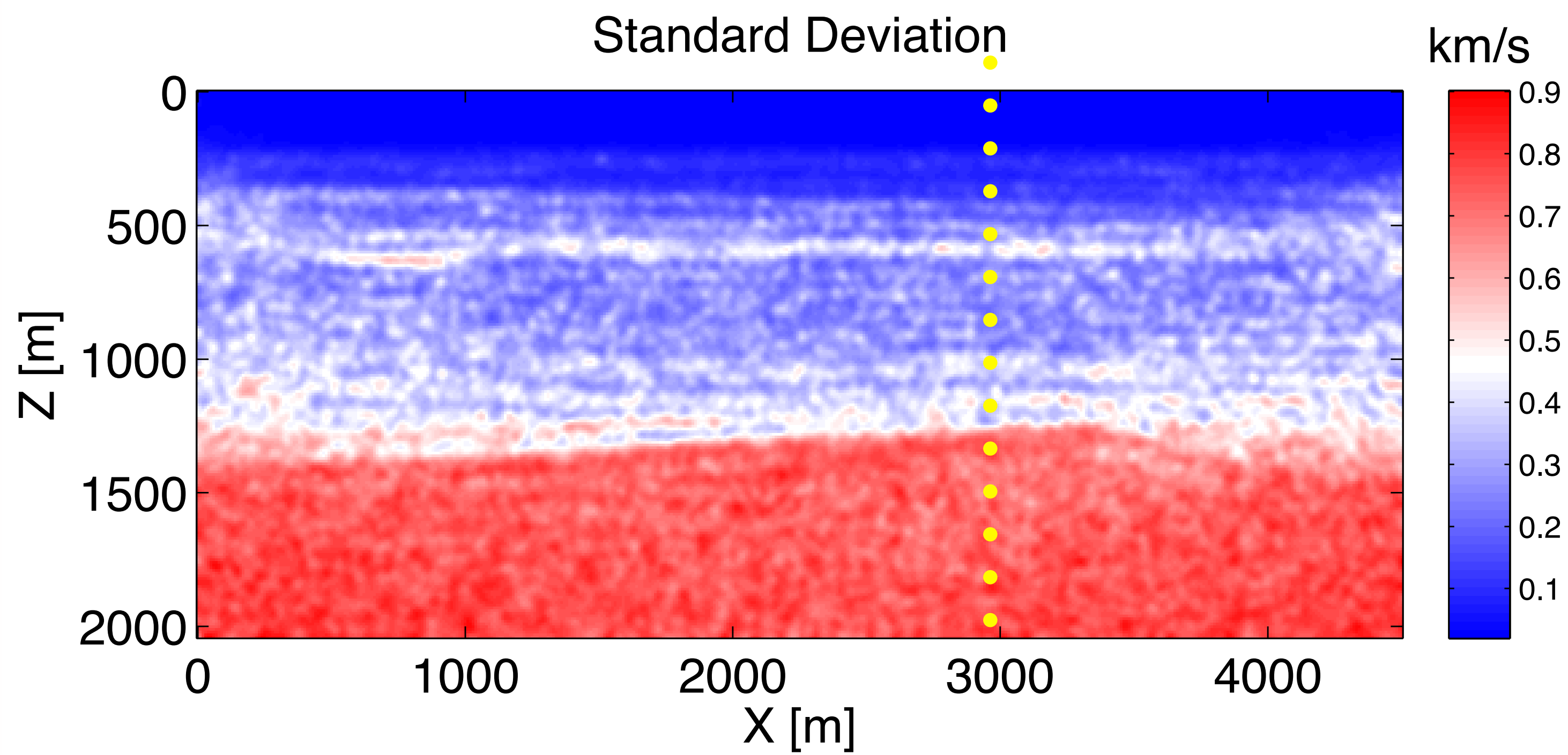
x = 1490m



Standard deviation



x = 2240m



Conclusions

- Using the L-BFGS Hessian, we reduce the computational cost of estimating the Hessian.
- Using the randomized source sub-sampling method, we decrease the computational cost of the McMC.

Future Work

- Use penalty method frame work to reset up the probability distribution function.
- Use different sampling method to analyze the posterior distribution.
- Uncertainty quantification for 3D FWI.

Acknowledgements

Thank you for your attention !

<https://www.slim.eos.ubc.ca/>



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