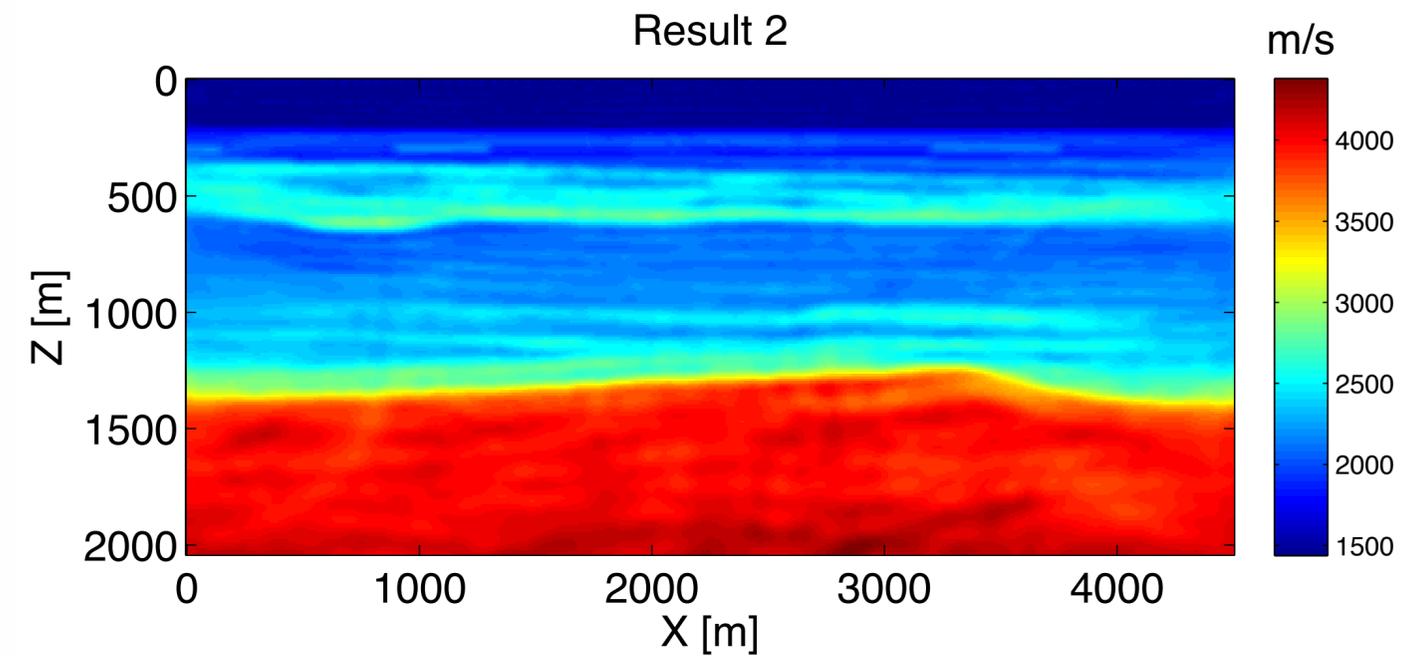
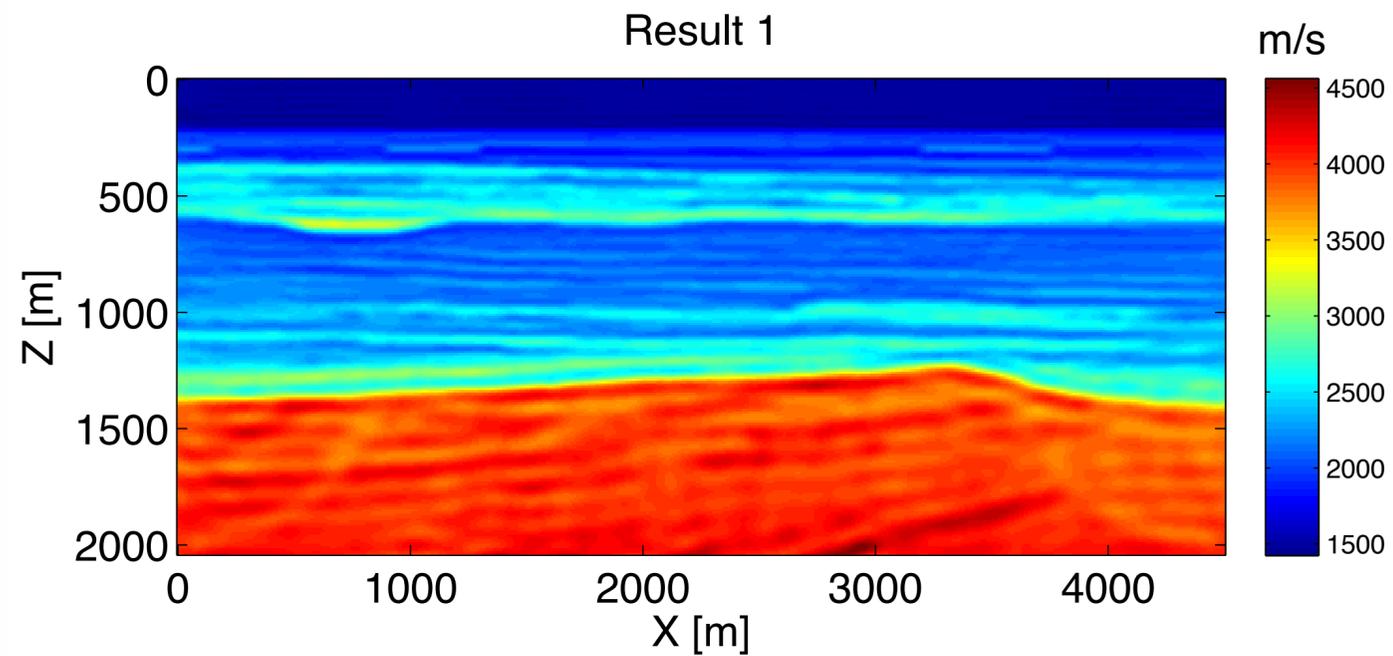


# A stochastic quasi-Newton MCMC method for uncertainty quantification of full-waveform inversion

Zhilong Fang, Felix J. Herrmann and Chia Ying Lee

# Motivation



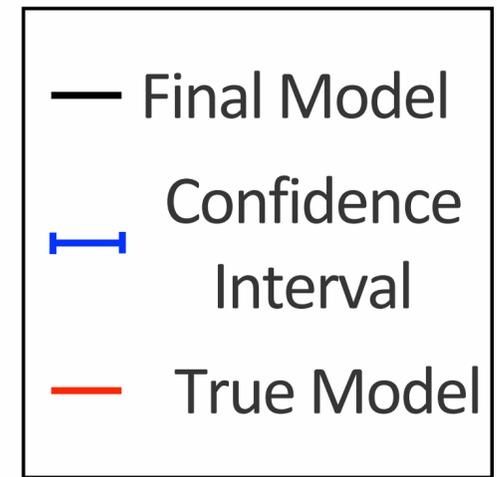
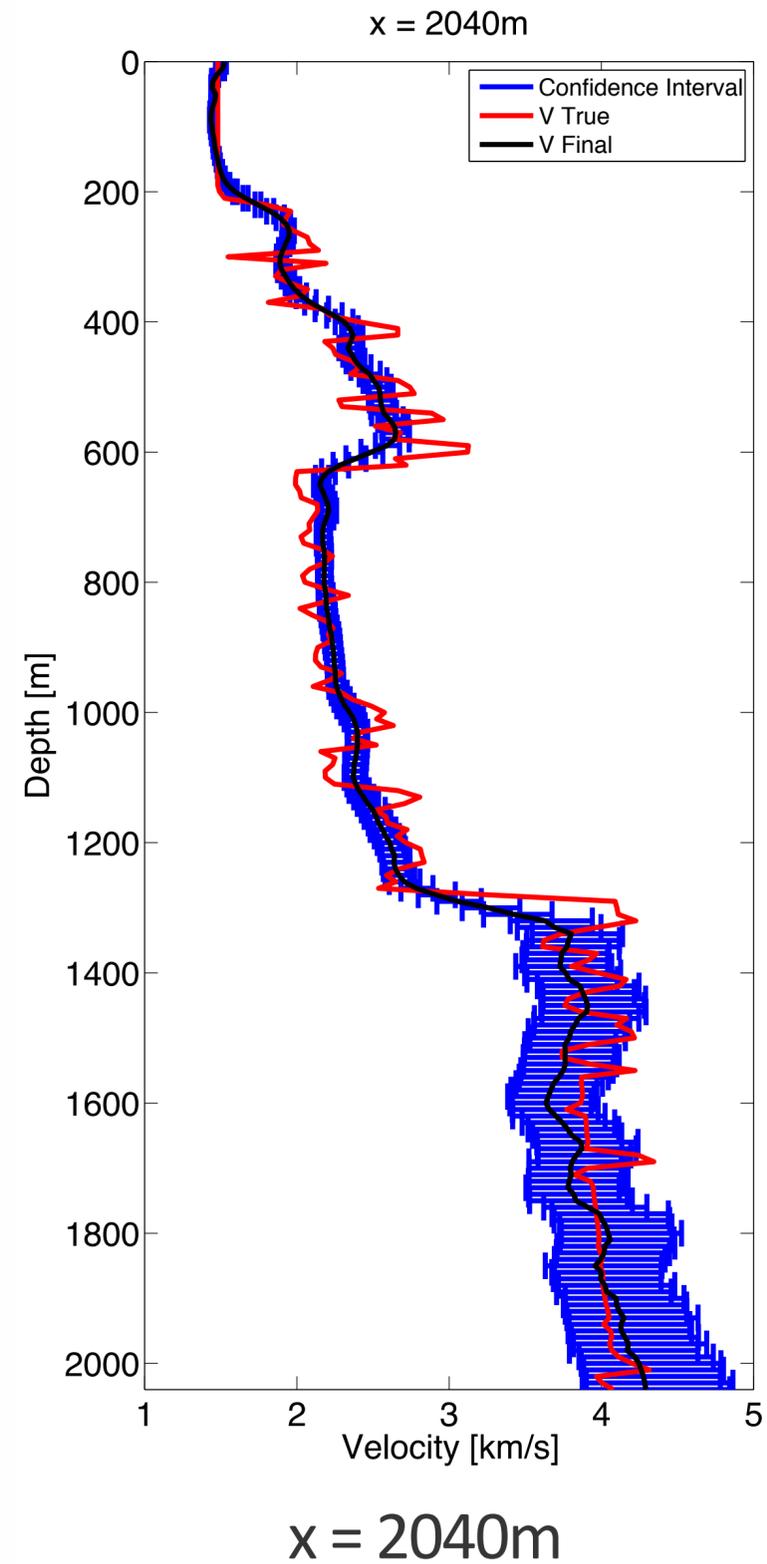
Which one is better?

Misfit?  
Eyeball norm?

Uncertainty?  
Standard Deviation?

# Confidence interval

Fast Accurate!!



## Bayesian theory

*Deterministic* inverse problem:

$$\mathbf{m}^* = \arg \min \left( \frac{1}{2N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{i\text{obs}}\|_{W_i}^2 + \frac{1}{2} \|\mathbf{m} - \bar{\mathbf{m}}\|_R^2 \right)$$

*Statistical* inverse problem with Bayesian theory:

$$\pi_{post}(\mathbf{m}) := \pi(\mathbf{m} | \mathbf{d}_{obs}) \propto \pi_{prior}(\mathbf{m}) \pi(\mathbf{d}_{obs} | \mathbf{m})$$

where  $\mathbf{m}$  is the model parameter, and  $\mathbf{d}_{obs}$  is the observed data

## Bayesian theory

Assume:

$$\text{noise} \sim \mathcal{N}(0, \Gamma_{noise})$$

$$\text{prior model distribution} \sim \mathcal{N}(\mathbf{m}_{prior}, \Gamma_{prior}).$$

Negative log-posterior of the posterior pdf:

$$V(\mathbf{m}) := -\log \pi_{post}(\mathbf{m}) := \frac{1}{2N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{inoise}^{-1}}^2 + \frac{1}{2} \|\mathbf{m} - \mathbf{m}_{prior}\|_{\Gamma_{prior}^{-1}}^2$$

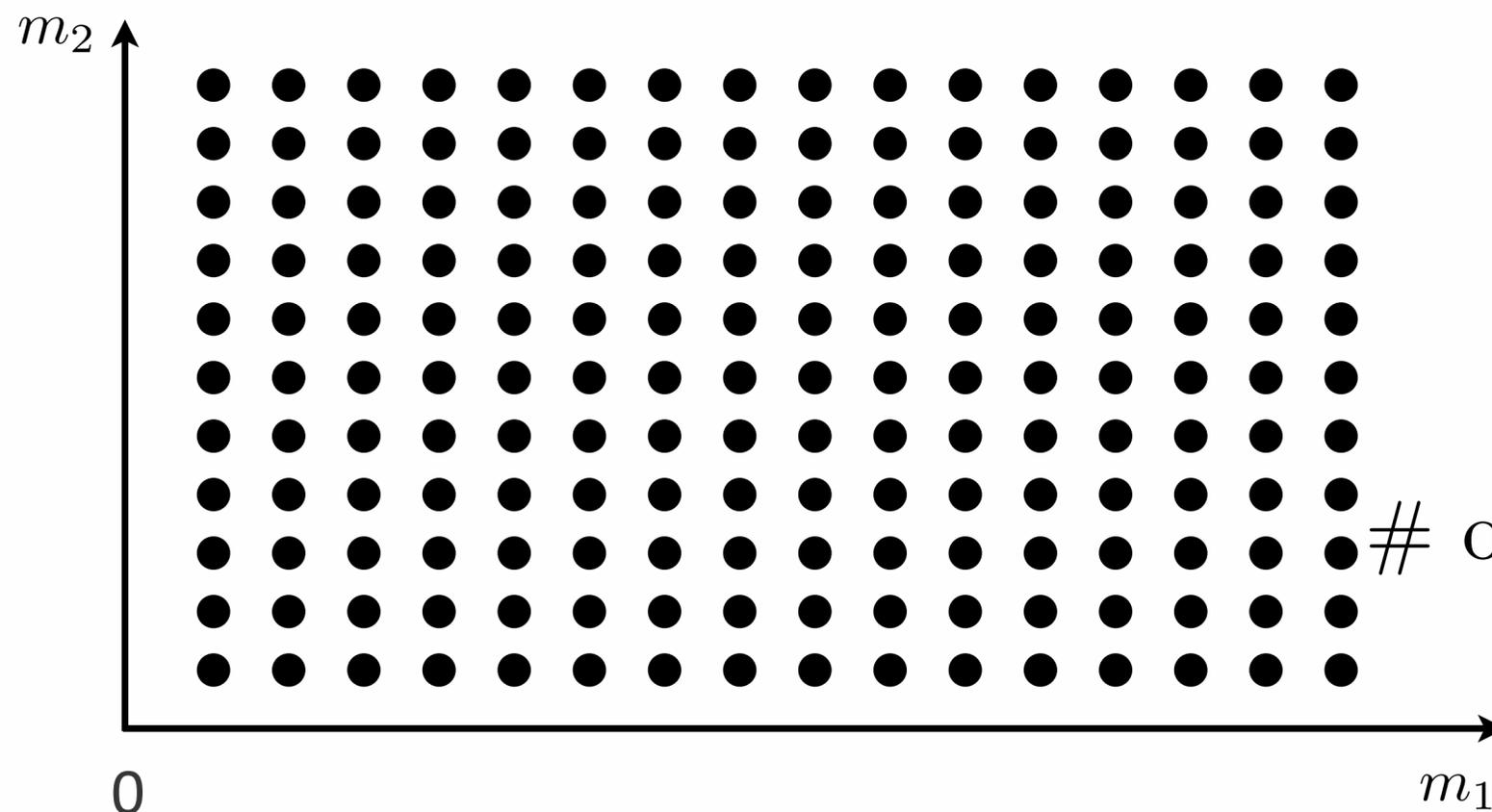
$$\mathbf{m}_{MAP} = \mathbf{m}^*$$

$$\|\mathbf{x}\|_{\Gamma_{noise}^{-1}}^2 := \mathbf{x}^T \Gamma_{noise}^{-1} \mathbf{x}$$

## Sampling the posterior pdf

How to obtain the posterior probability density function?

- Compute  $\pi_{post}(\mathbf{m})$  by discretization?  $N = N_{sub\_sample} * N_{para}$



$$\begin{aligned} N_{para} &>> 10^6 \\ N_{sub\_sample} &>> 10^2 \\ N &>> 10^8 \\ N_s &>> 10^2 \end{aligned}$$

# of PDE solvers  $>> N * N_s$

## Approximate the pdf

How to obtain the posterior probability density function?

- Markov chain Monte Carlo method ? Metropolis - Hasting method

At sample  $\mathbf{m}_k$

Draw sample  $\mathbf{y}$  from the proposal distribution  $\tilde{\pi}_k(\mathbf{m})$

if  $\min\left(1, \frac{\pi_{post}(\mathbf{y})\tilde{\pi}_y(\mathbf{m}_k)}{\pi_{post}(\mathbf{m}_k)\tilde{\pi}_k(\mathbf{y})}\right) > \alpha$

    set  $\mathbf{m}_{k+1} = \mathbf{y}$

else

    regenerate  $\mathbf{y}$

end

## Approximate the pdf

How to obtain the posterior probability density function?

- Stochastic Newton MCMC?

$$V(\mathbf{m}) := -\log \pi_{post}(\mathbf{m}) := \frac{1}{2N_s} \sum_{i=1}^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 + \frac{1}{2} \|\mathbf{m} - \mathbf{m}_{prior}\|_{\Gamma_{prior}^{-1}}^2$$

$$\tilde{\pi}_k(\mathbf{m}) \sim \mathcal{N}(\mathbf{m}_k - \mathbf{H}_k^{-1} \mathbf{g}_k, \mathbf{H}_k^{-1})$$

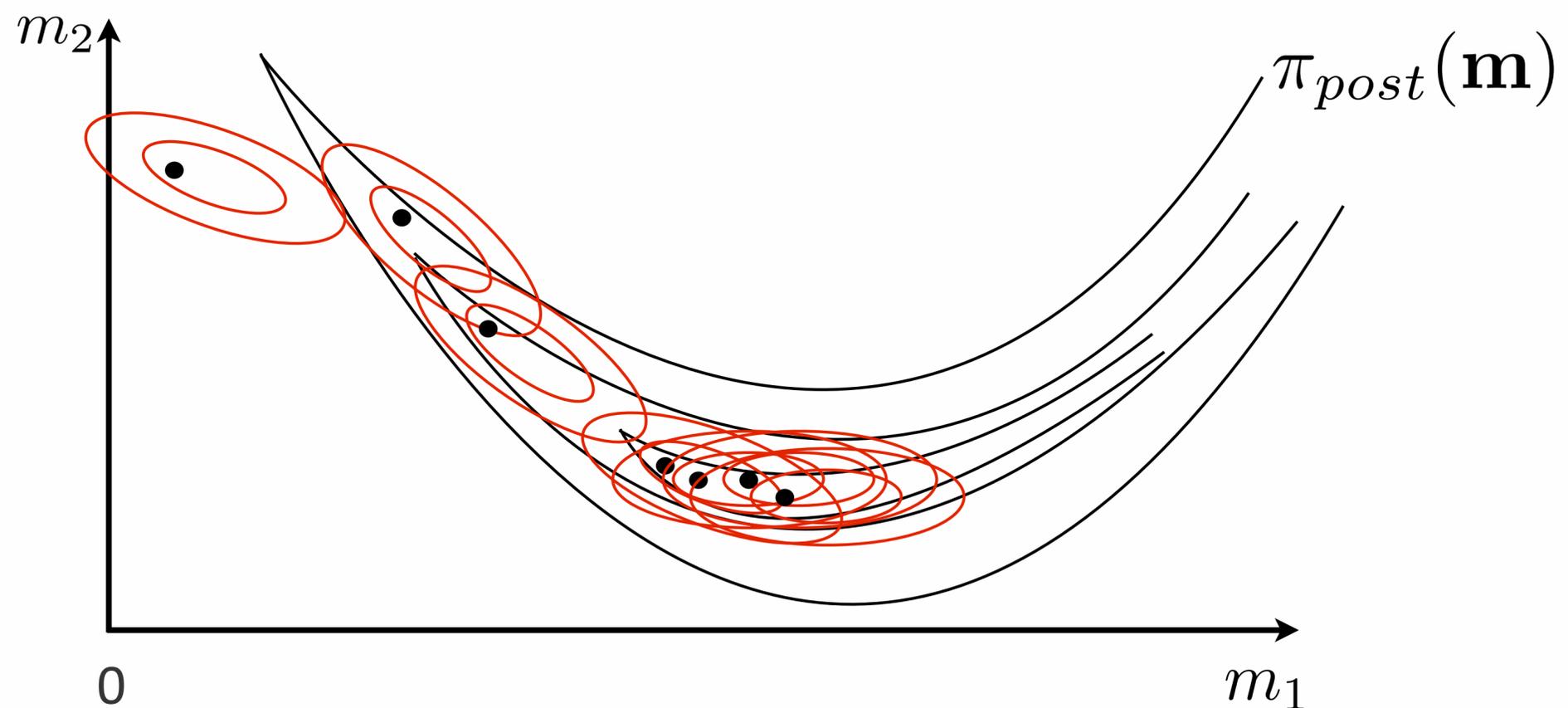
$$\begin{bmatrix} \mathbf{L}^T \\ \mathbf{L} \end{bmatrix} \begin{bmatrix} \mathbf{H}_{misfit} \\ \mathbf{L} \end{bmatrix} \approx \begin{bmatrix} \mathbf{V}_r \\ \mathbf{D}_r \\ \mathbf{V}_r^T \end{bmatrix}$$

$n \times n$       $n \times n$       $n \times n$       $n \times r$       $r \times r$       $r \times n$

$r \ll n$   
(James Martin et al, 2012)

## Approximate the pdf

$$\tilde{\pi}_k(\mathbf{m}) \sim \mathcal{N}(\mathbf{m}_k - \mathbf{H}_k^{-1} \mathbf{g}_k, \mathbf{H}_k^{-1})$$



## Approximate the pdf

### Challenges:

- Low - rank approximation may not be correct.
- The computational cost of estimating Hessian is huge.
- The computational cost of calculating the posterior distribution probability density function is huge.

Computational Cost  $\sim \mathcal{O}(N_{\text{sample}} * N_s)$

## BFGS Hessian

At the stage  $k$ : approximated Hessian  $\mathbf{B}_k$  and gradient  $\mathbf{g}_k$ ,  
Update  $\mathbf{B}_k$  by adding two rank one matrices:

$$\mathbf{B}_{k+1} = \mathbf{B}_k + \mathbf{U}_k + \mathbf{V}_k$$

where  $\mathbf{B}_{k+1}$  should satisfy

$$\mathbf{B}_{k+1}(\mathbf{x}_{k+1} - \mathbf{x}_k) = \mathbf{g}_{k+1} - \mathbf{g}_k$$

No additional computational cost!!!



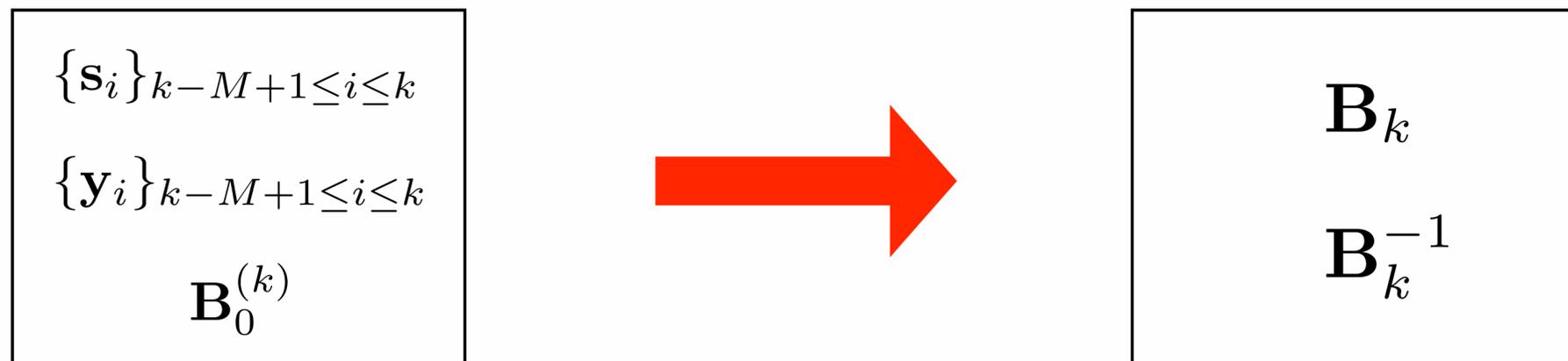
$$\begin{aligned} \mathbf{y}_k &= \mathbf{g}_{k+1} - \mathbf{g}_k \\ \mathbf{s}_k &= \mathbf{x}_{k+1} - \mathbf{x}_k, \end{aligned}$$

$$\mathbf{B}_{k+1} = \mathbf{B}_k + \frac{\mathbf{y}_k \mathbf{y}_k^T}{\mathbf{y}_k^T \mathbf{s}_k} - \frac{\mathbf{B}_k \mathbf{s}_k \mathbf{s}_k^T \mathbf{B}_k}{\mathbf{s}_k^T \mathbf{B}_k \mathbf{s}_k}$$

$$\mathbf{B}_{k+1}^{-1} = \left( \mathbf{I} - \frac{\mathbf{s}_k \mathbf{y}_k^T}{\mathbf{y}_k^T \mathbf{s}_k} \right) \mathbf{B}_k^{-1} \left( \mathbf{I} - \frac{\mathbf{y}_k \mathbf{s}_k^T}{\mathbf{y}_k^T \mathbf{s}_k} \right) + \frac{\mathbf{s}_k \mathbf{s}_k^T}{\mathbf{y}_k^T \mathbf{s}_k}$$

# I-BFGS Hessian with pseudo GN Hessian as a starter

Limited memory BFGS Hessian:



Choice of  $\mathbf{B}_0^{(k)}$ : pseudo GN Hessian

## Randomized source subsampling

To speed up the computation of posterior pdf:

$$V(\mathbf{m}) := -\log \pi_{post}(\mathbf{m}) := \frac{1}{2N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 + \frac{1}{2} \|\mathbf{m} - \mathbf{m}_{prior}\|_{\Gamma_{prior}^{-1}}^2$$

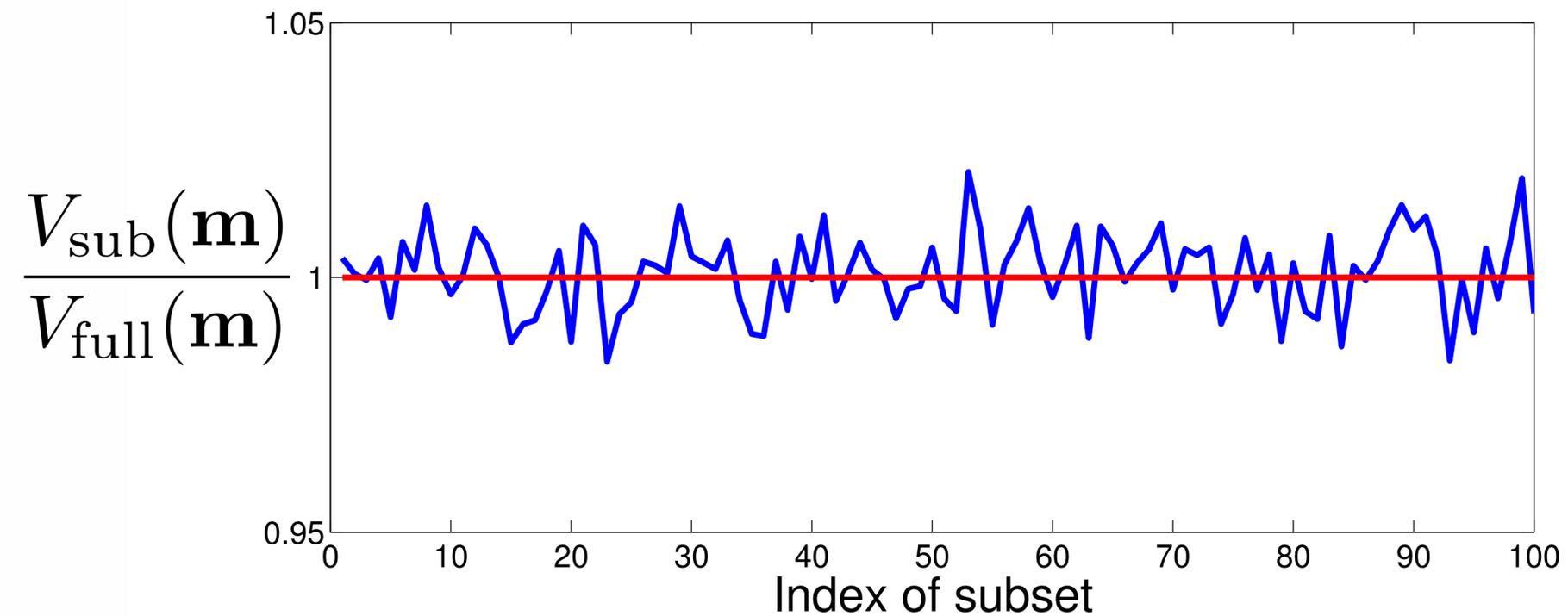
Randomized source subsampling:

$$\frac{1}{N_s} \sum_1^{N_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 = \frac{1}{\|\mathcal{I}_s\|} \sum_{i \in \mathcal{I}_s} \|f_i(\mathbf{m}) - \mathbf{d}_{iobs}\|_{\Gamma_{innoise}^{-1}}^2 + \epsilon$$

## Randomized source subsampling

Computational Cost:  $\mathcal{O}(N_s) \rightarrow \mathcal{O}(N_{rs}), N_{rs} \ll N_s$

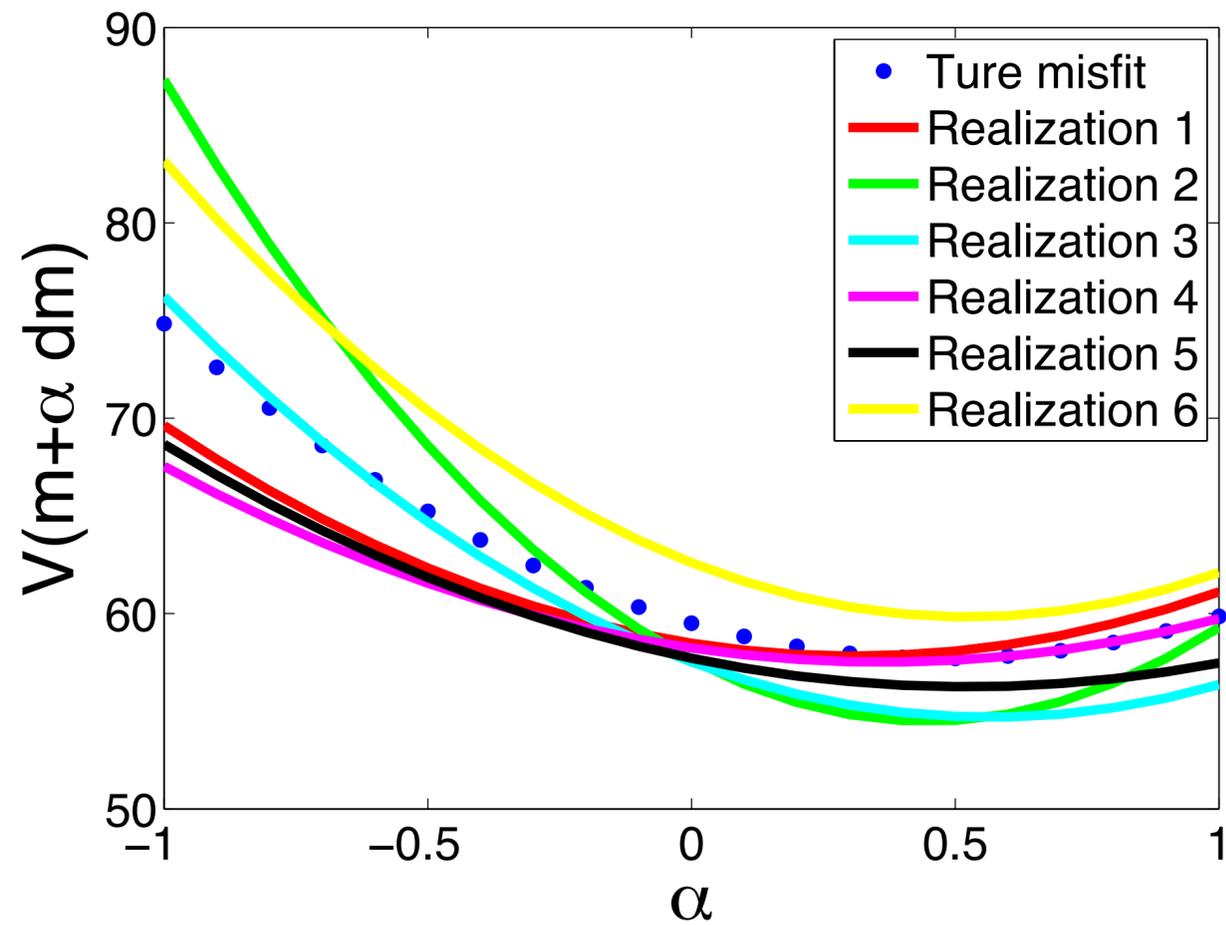
Subset misfit vs Full misfit



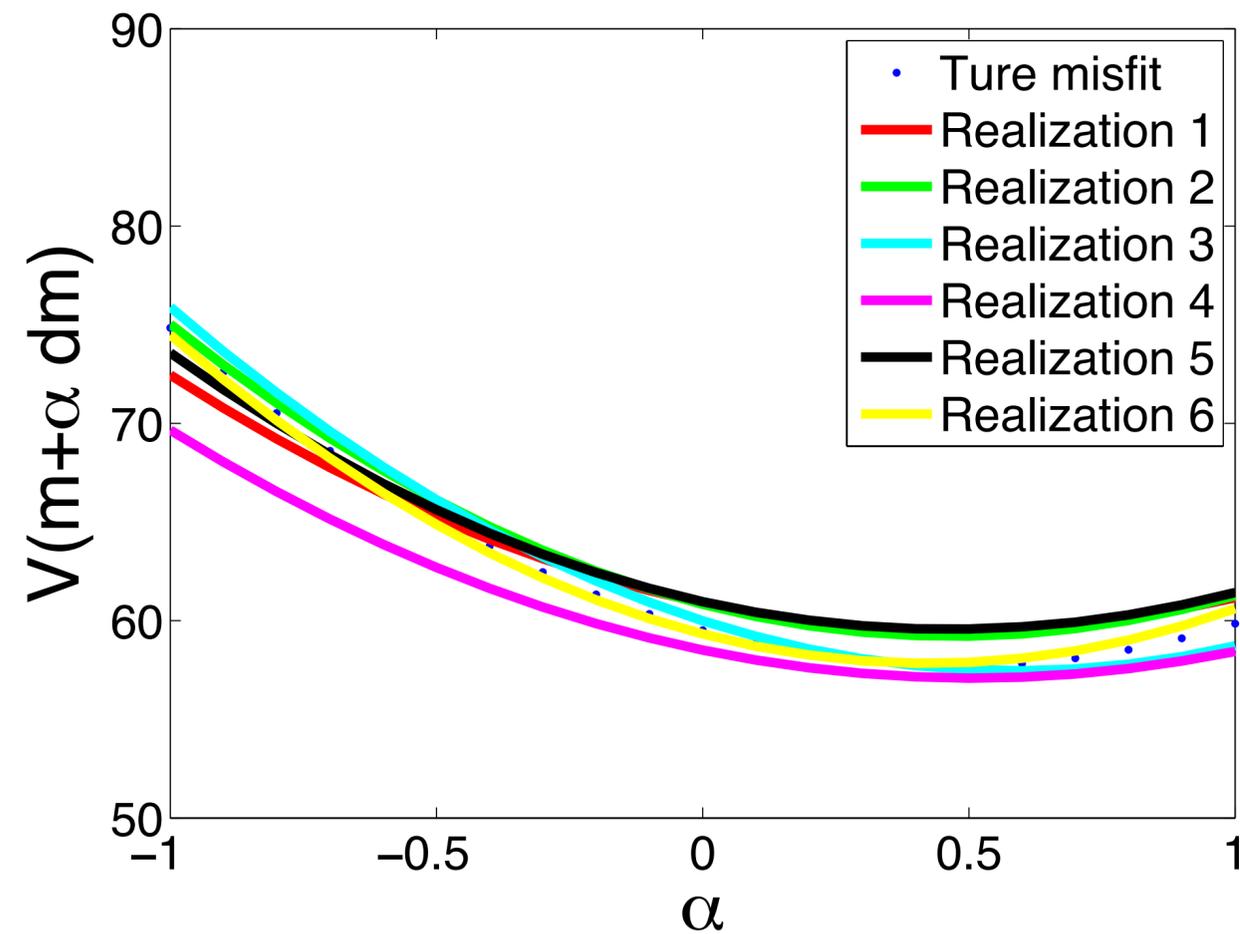
# Randomized source subsampling

Computational Cost:  $\mathcal{O}(N_s) \rightarrow \mathcal{O}(N_{rs})$

5 / 91 shots

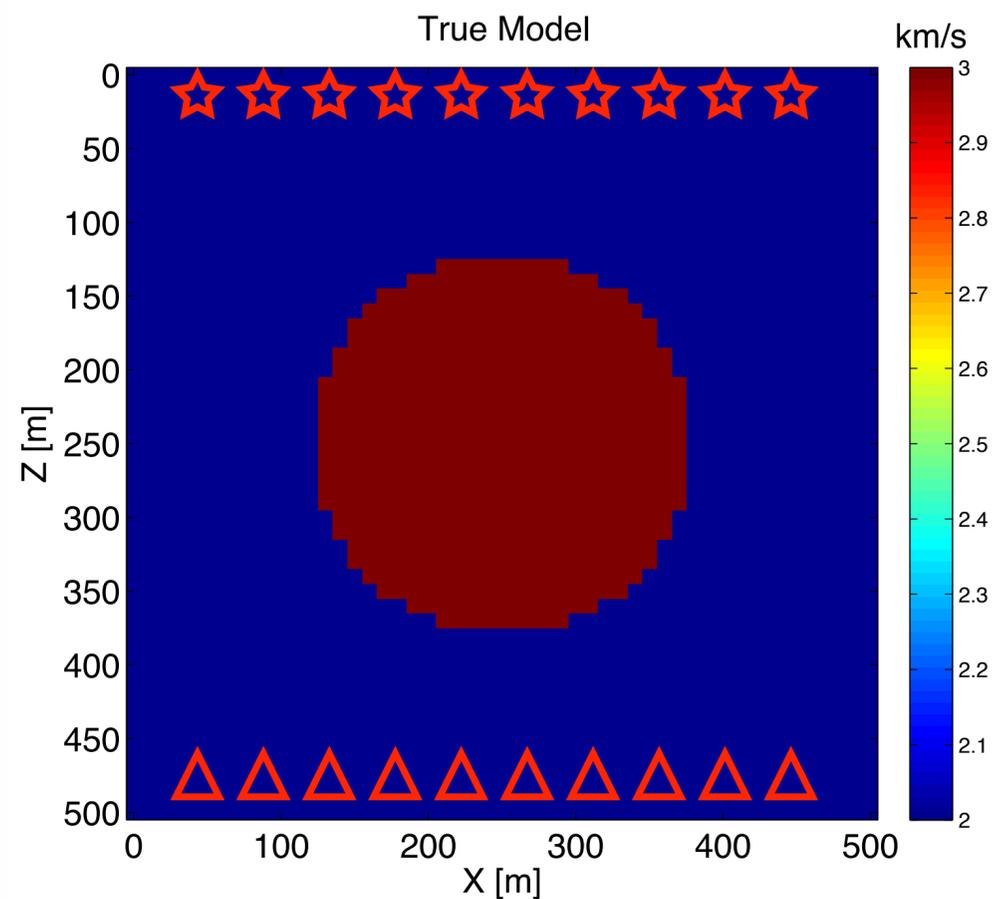


20 / 91 shots



# Numerical Experiments

## Camambert model



- ☆ - source
- △ - receiver

Statistical parameter to be inverted:

Standard deviation :  $\sigma$

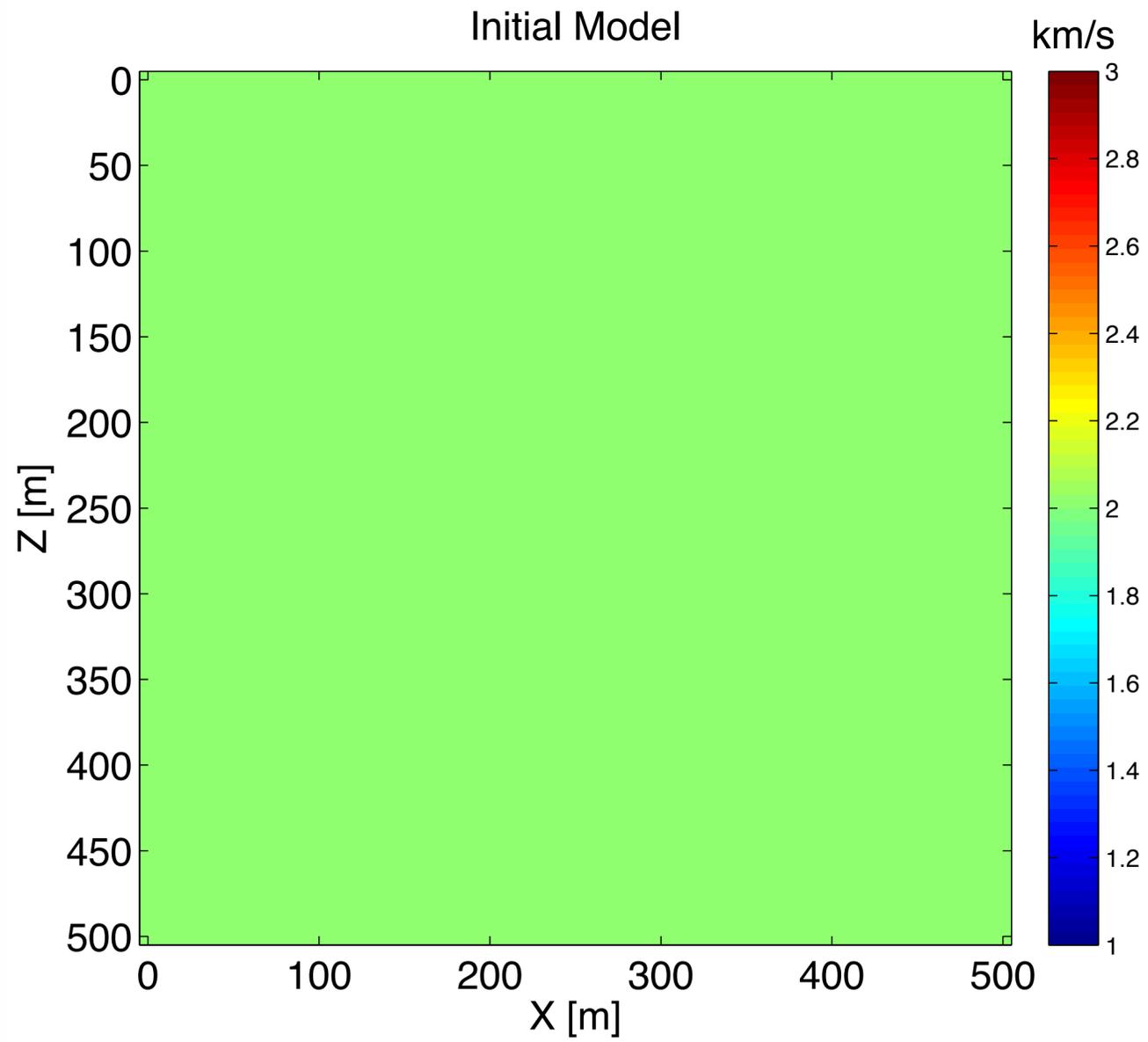
Confidence interval:  $P(\mathbf{m} \in I_{ci}) \geq \alpha$

Acquisition Geometry:

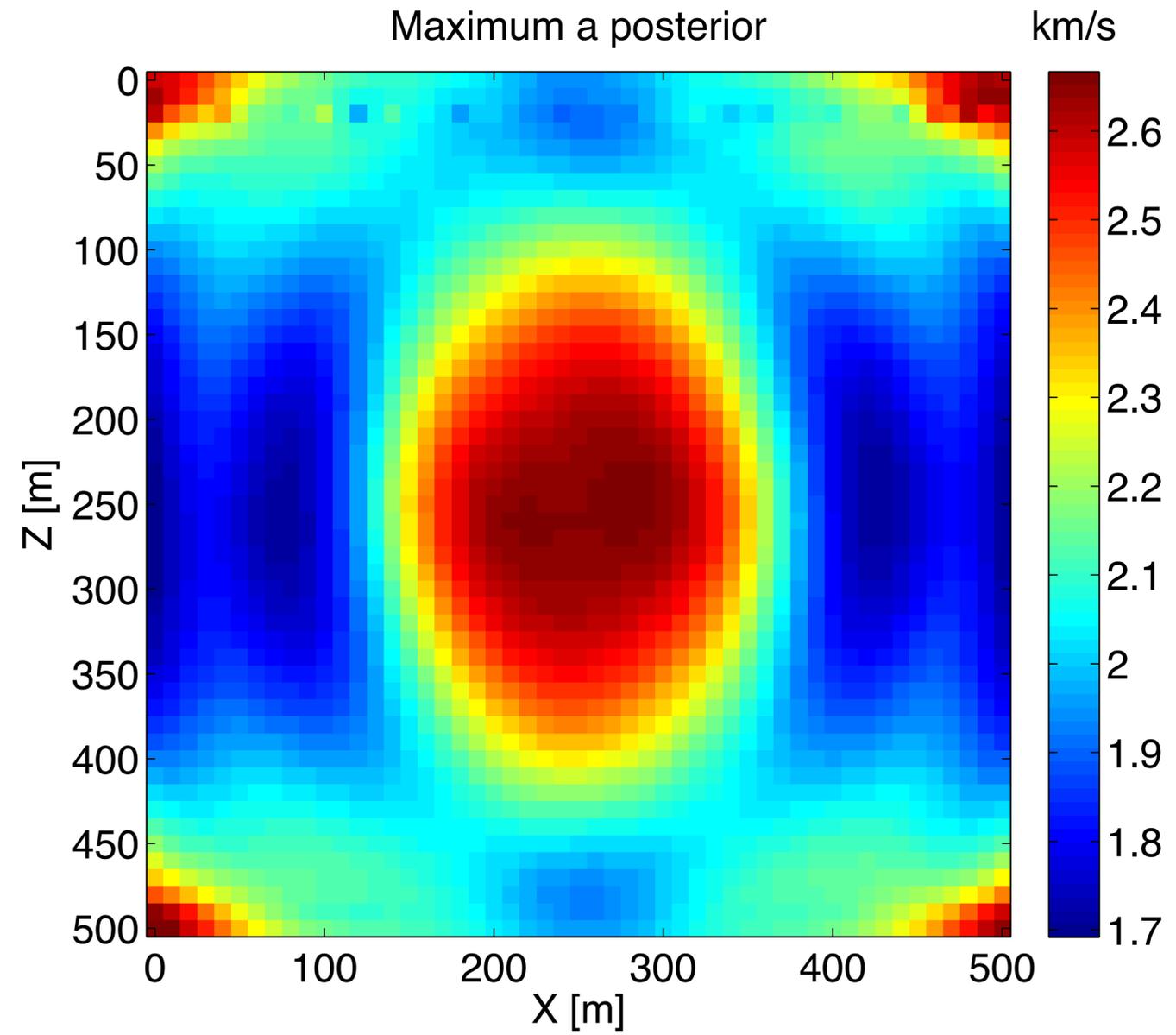
26 shots

51 receivers

10 frequencies

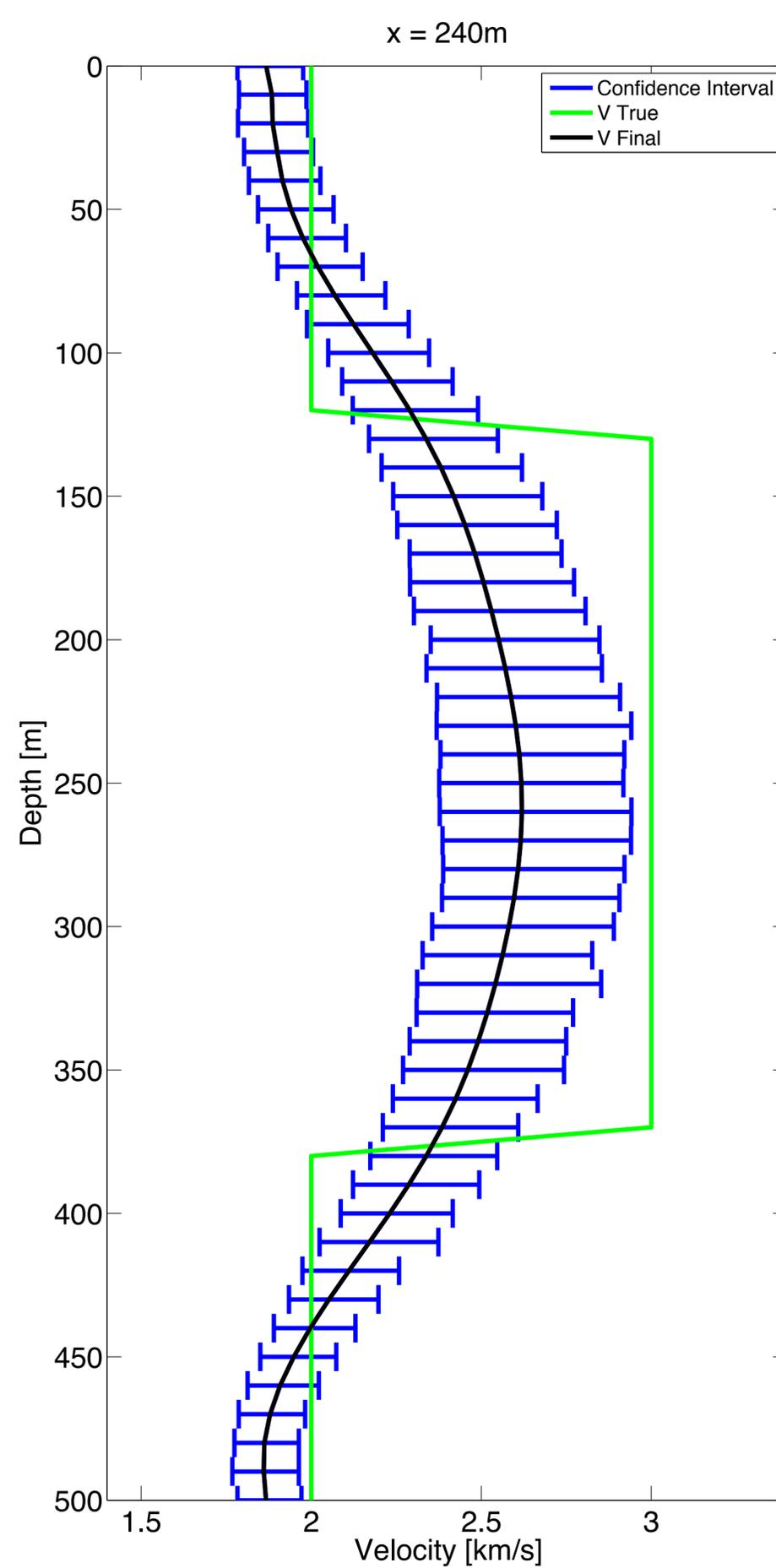
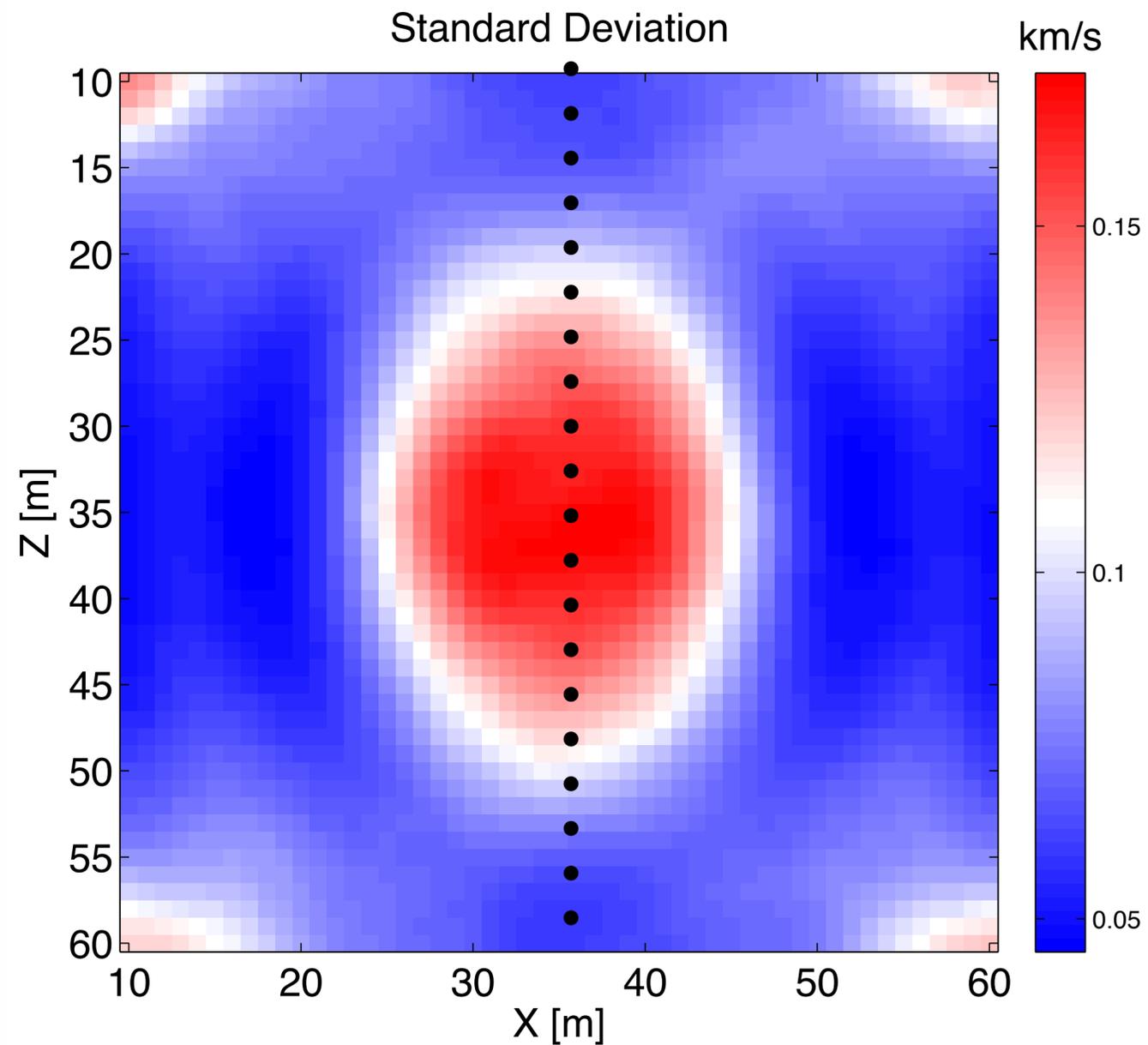


Initial model

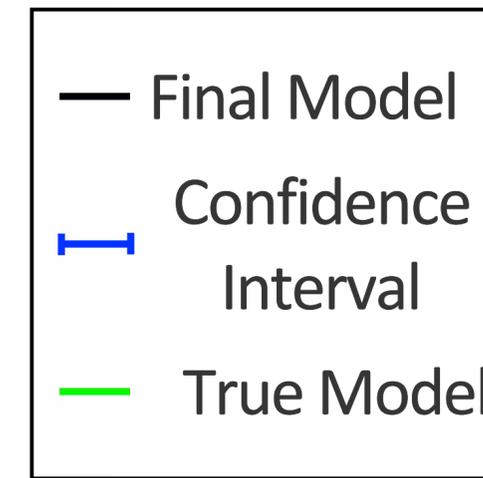


MAP

Nrs = 5

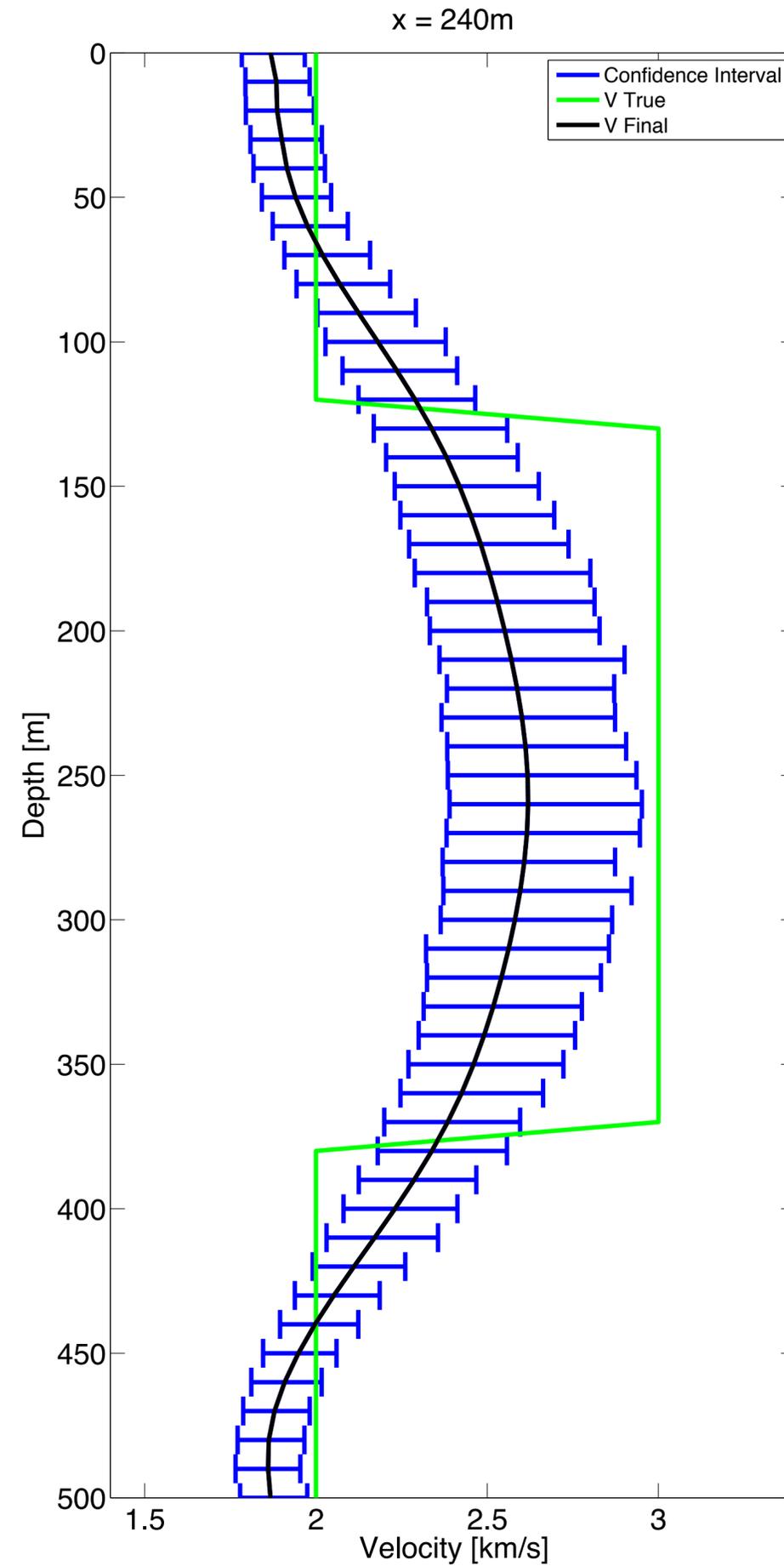
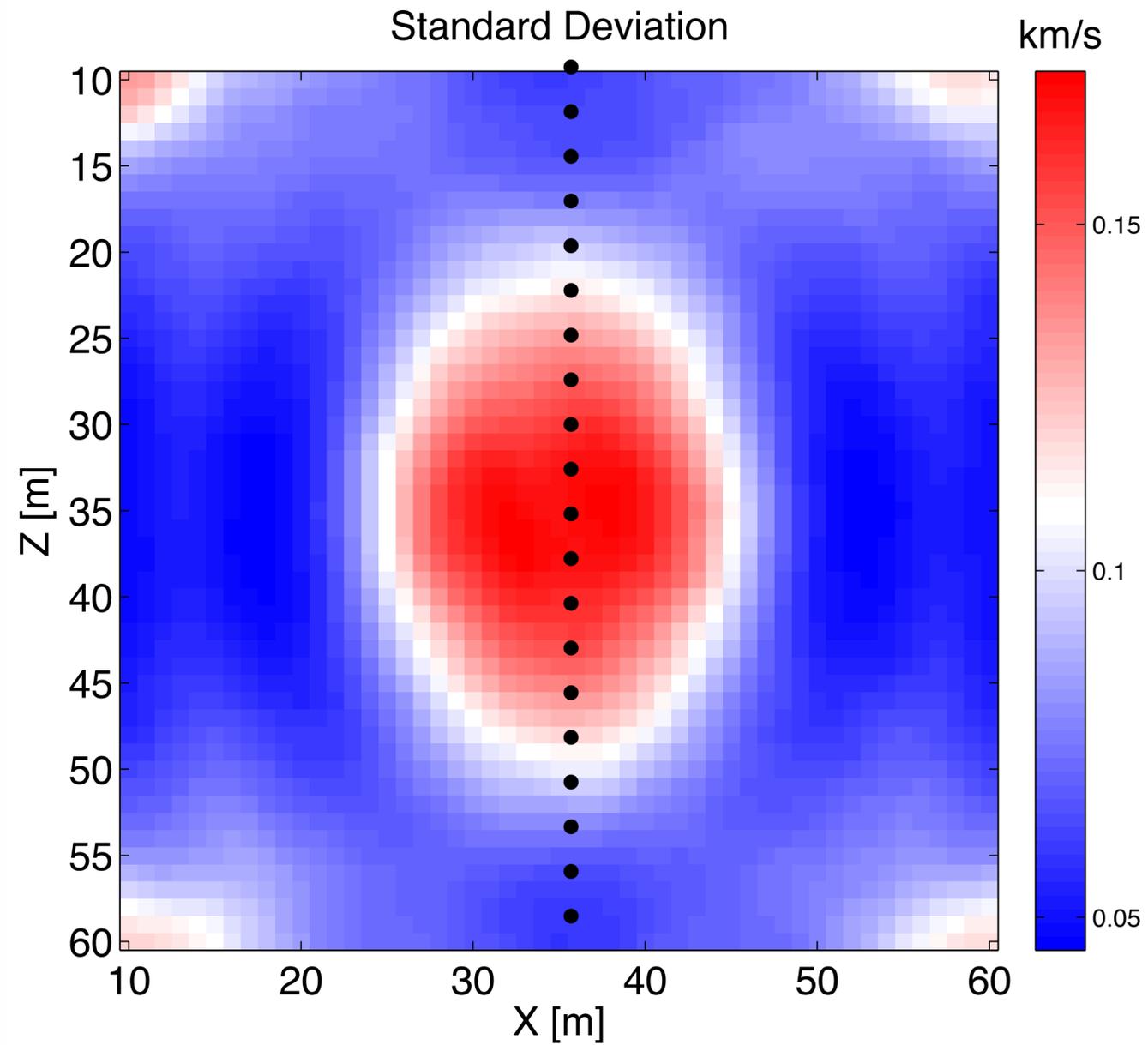


$\alpha = 0.9$



Confidence Interval

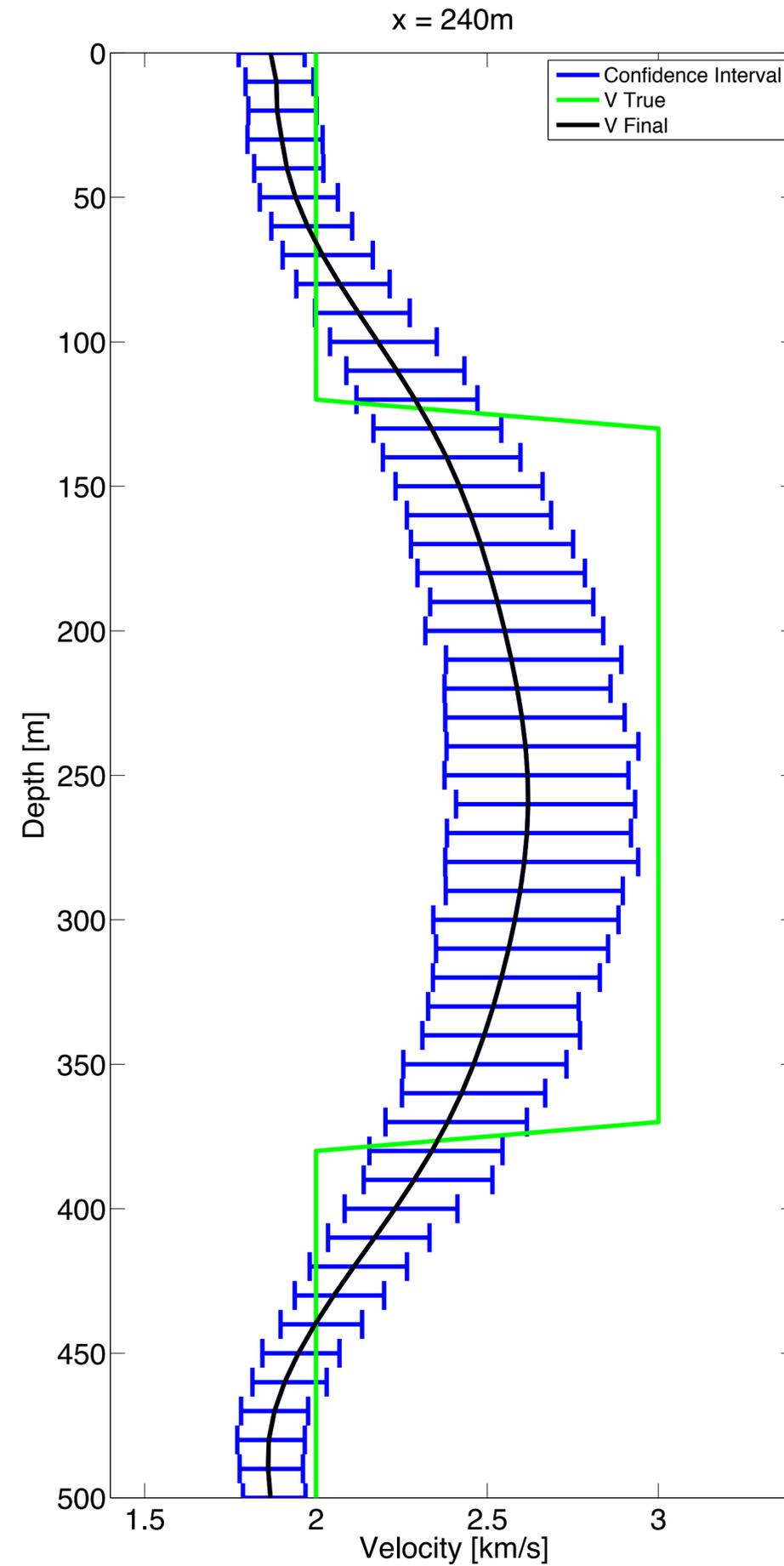
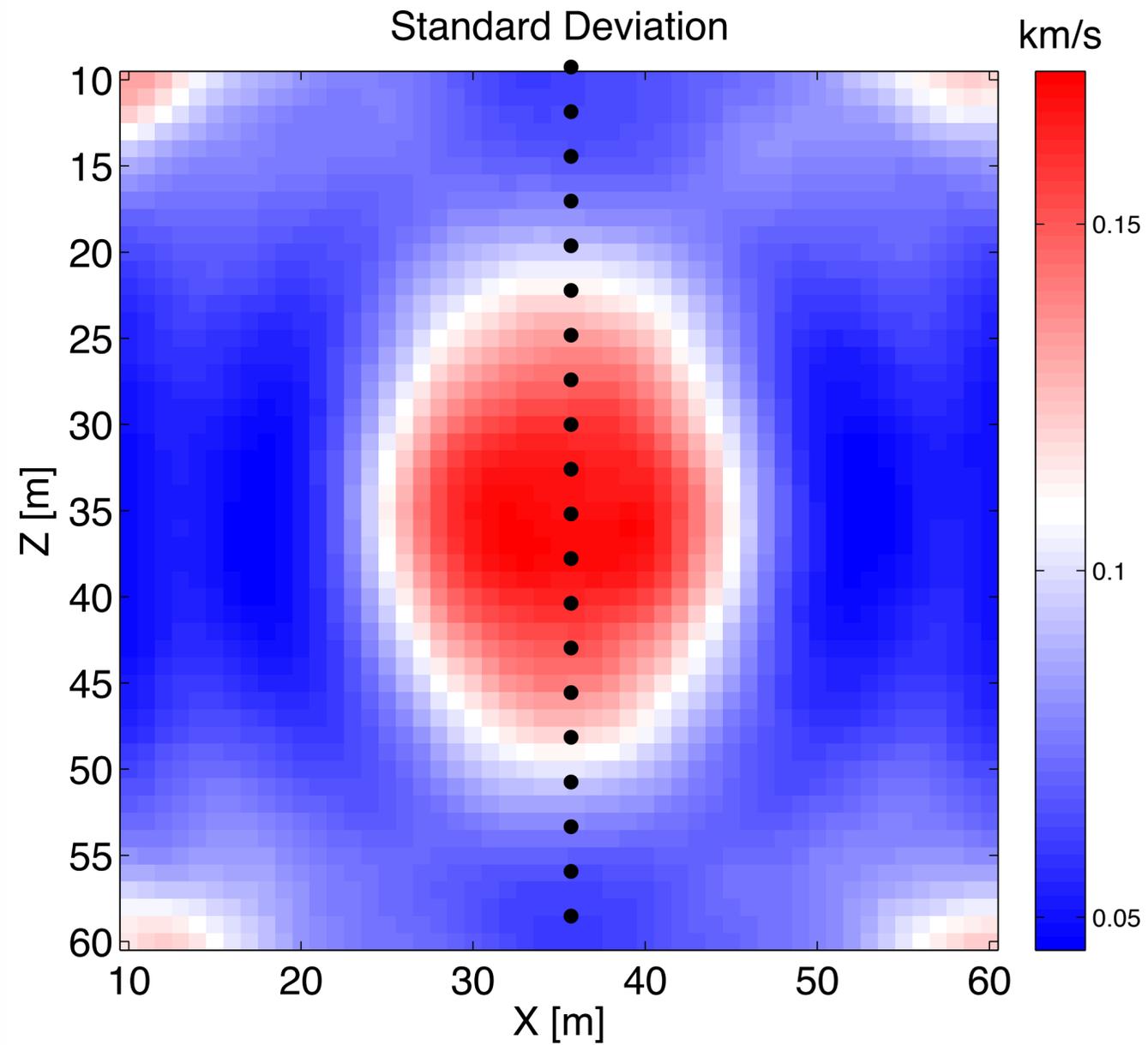
Nrs = 10



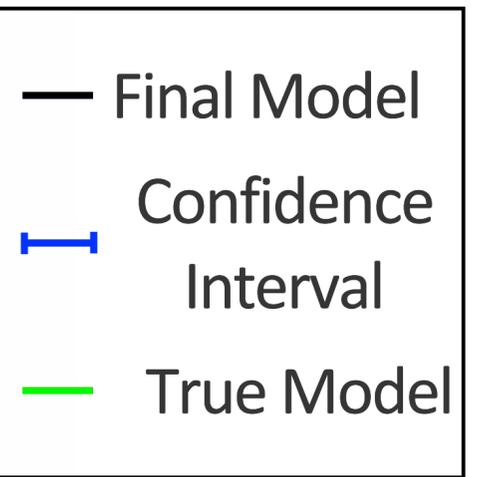
$\alpha = 0.9$

Confidence Interval

Nrs = 26



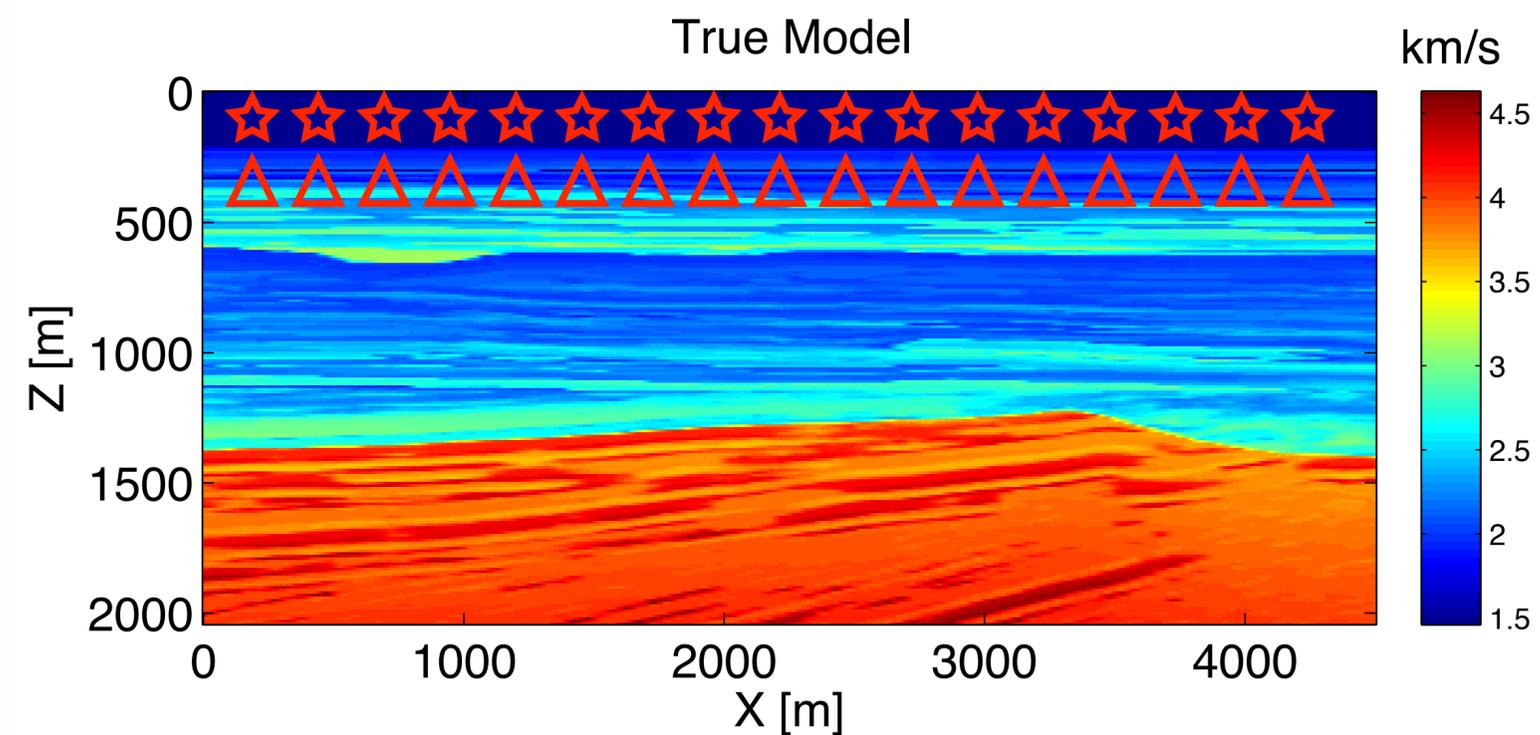
$\alpha = 0.9$



Confidence Interval

# Numerical Experiments

## BG model



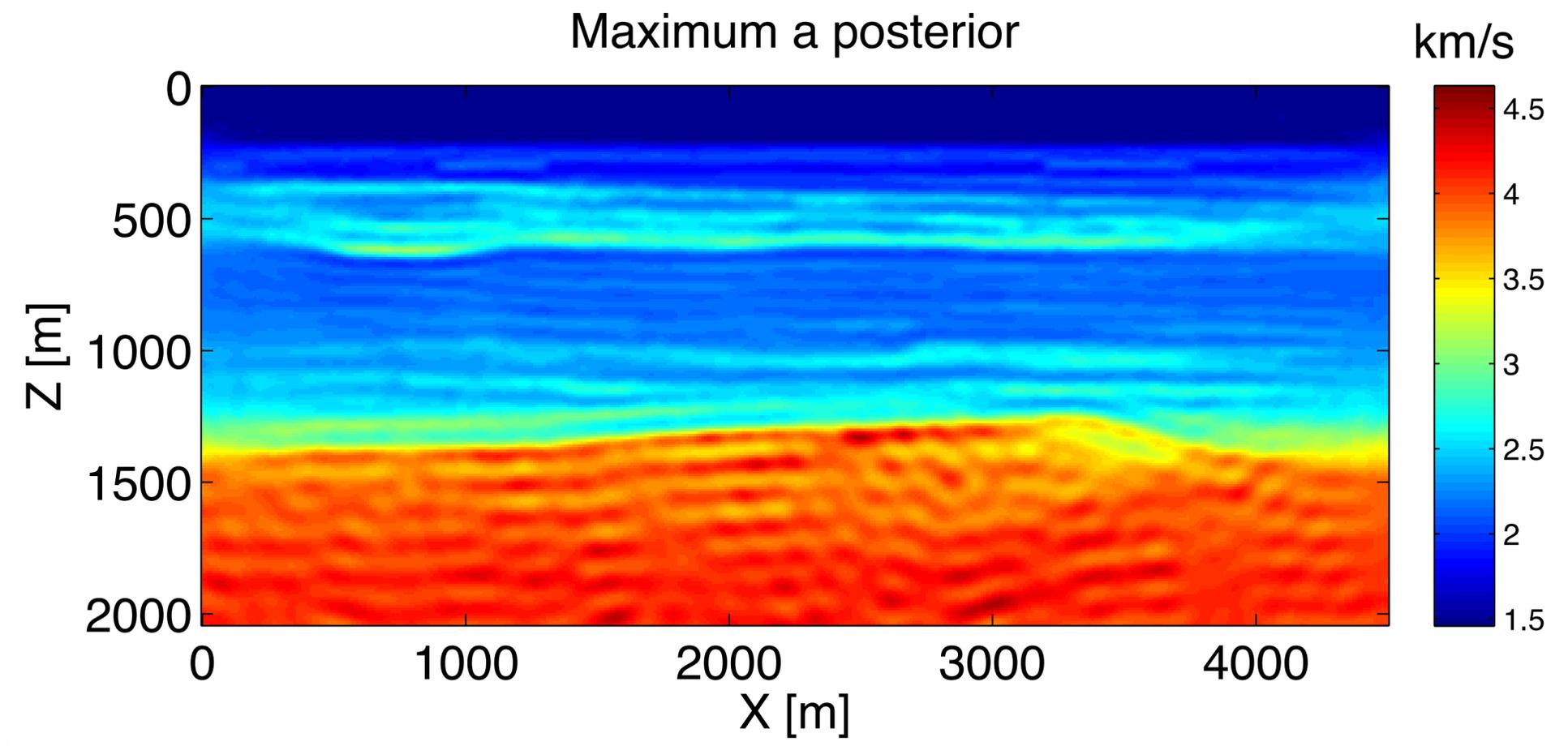
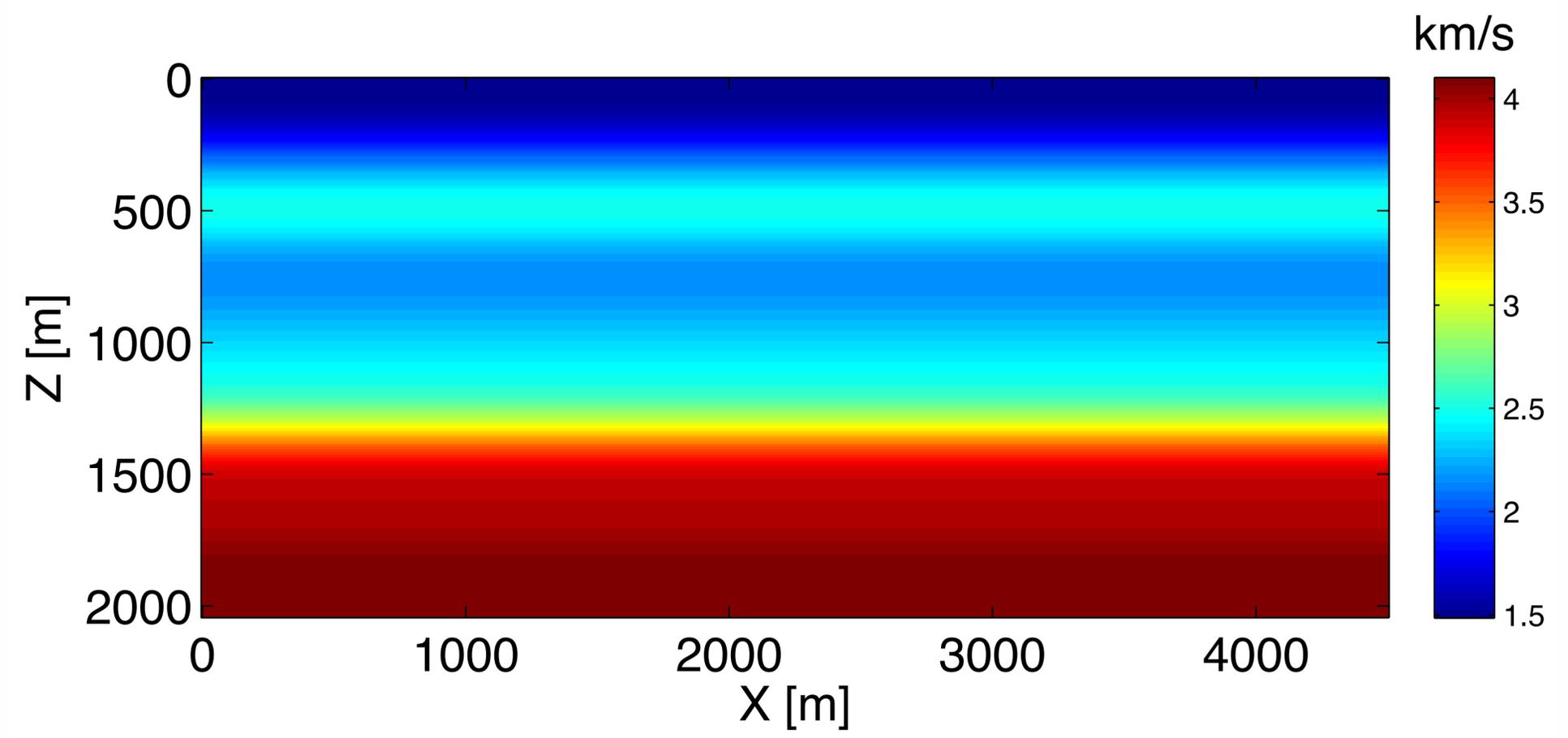
☆ - source  
△ - receiver

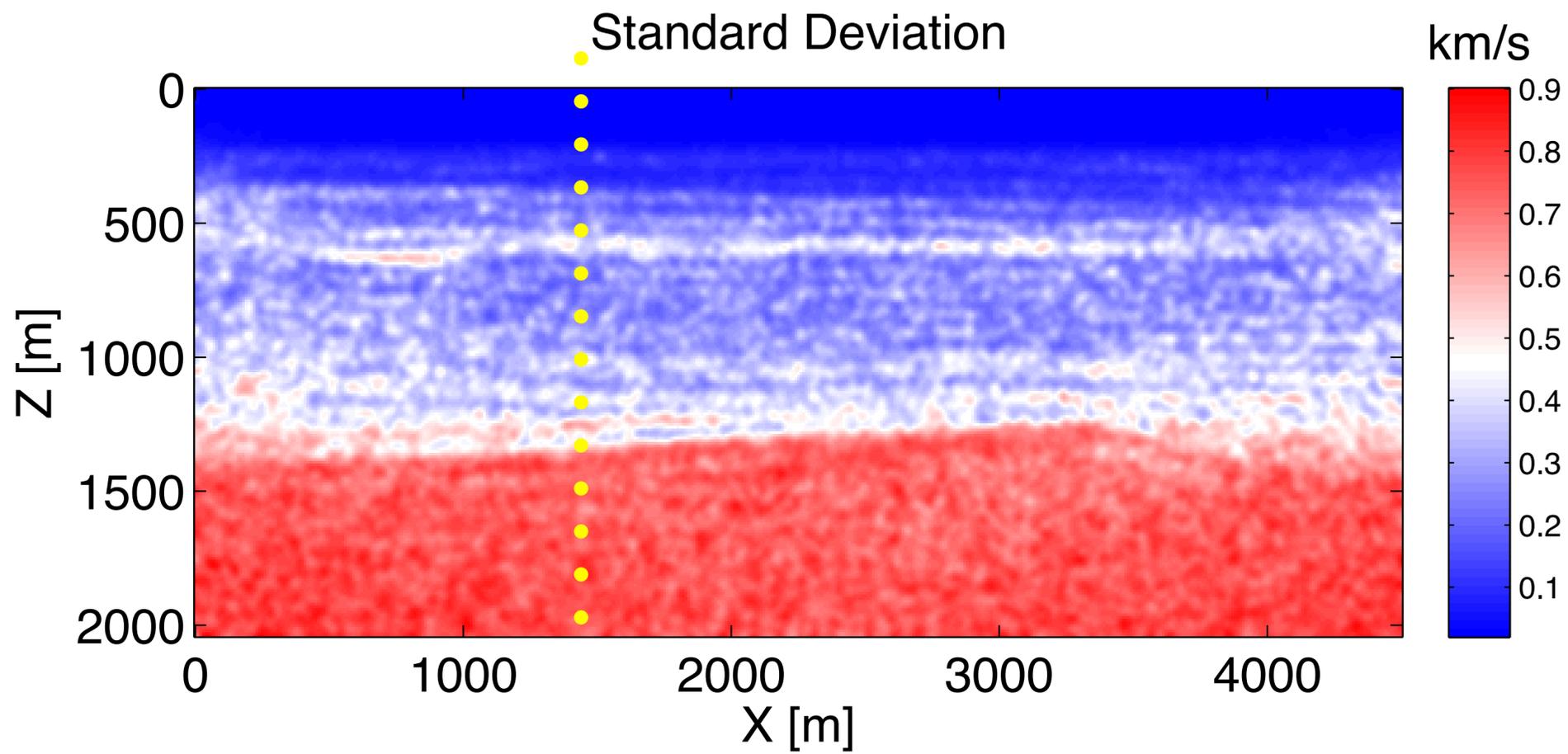
Acquisition Geometry:

91 shots

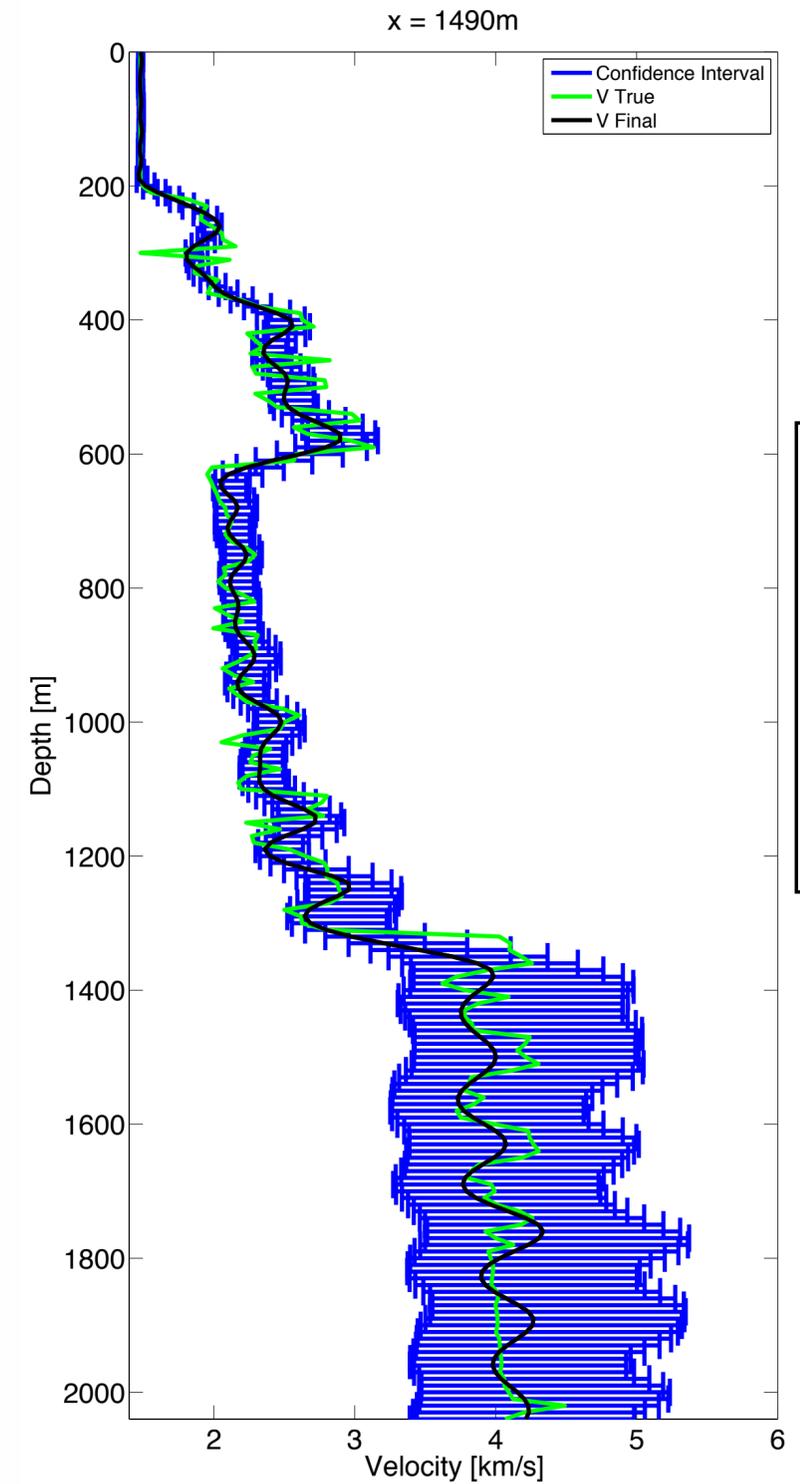
451 receivers

15 frequencies from 3Hz to 17Hz

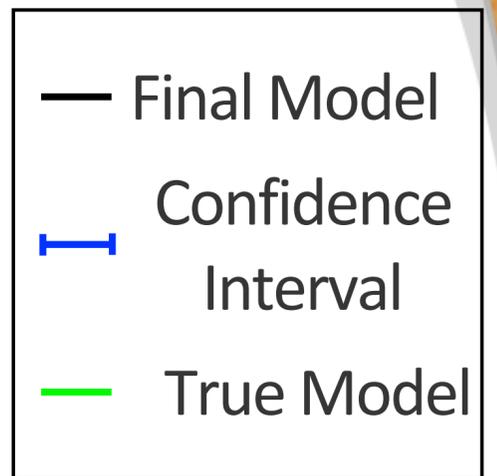


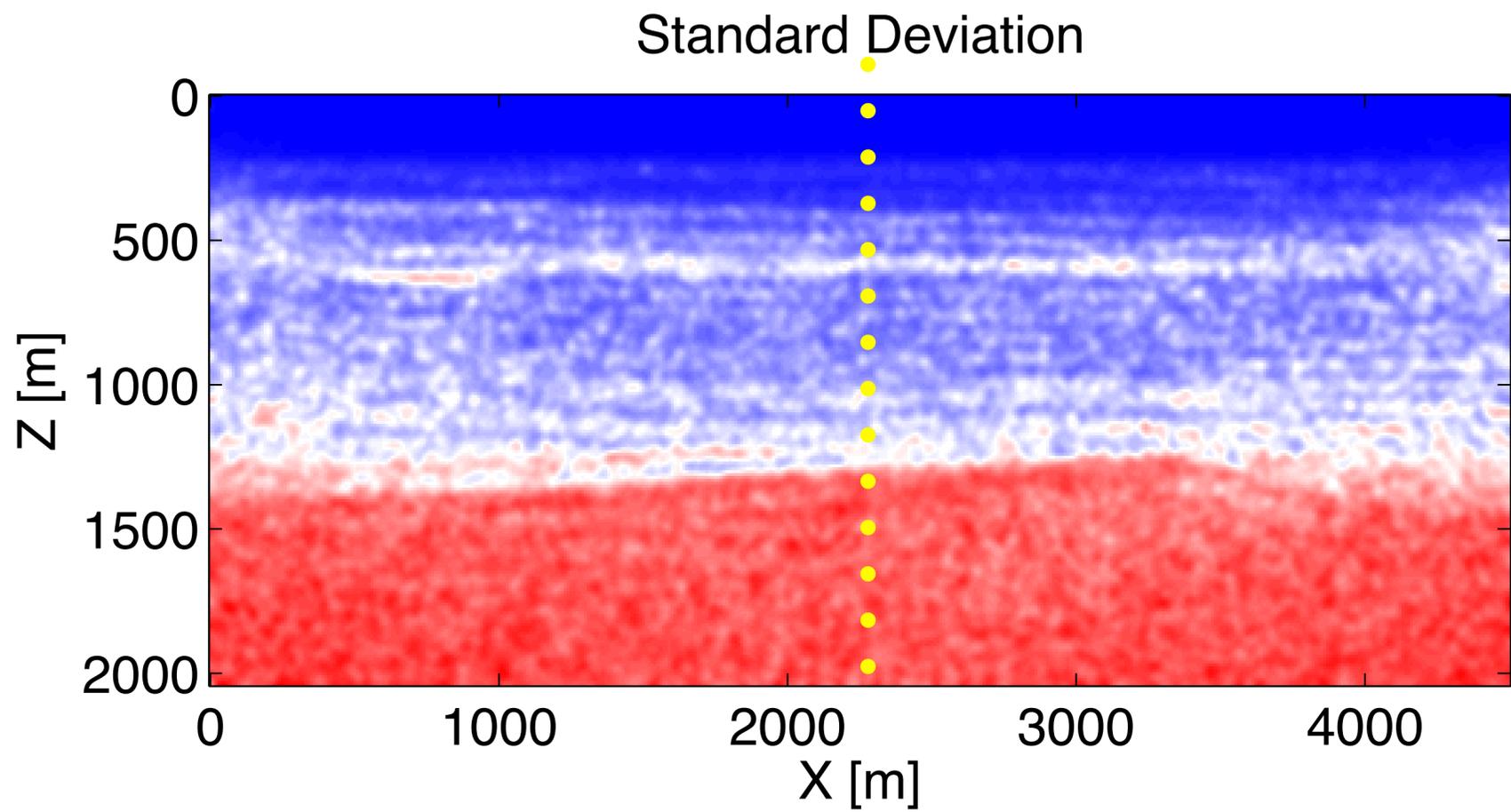


Standard deviation

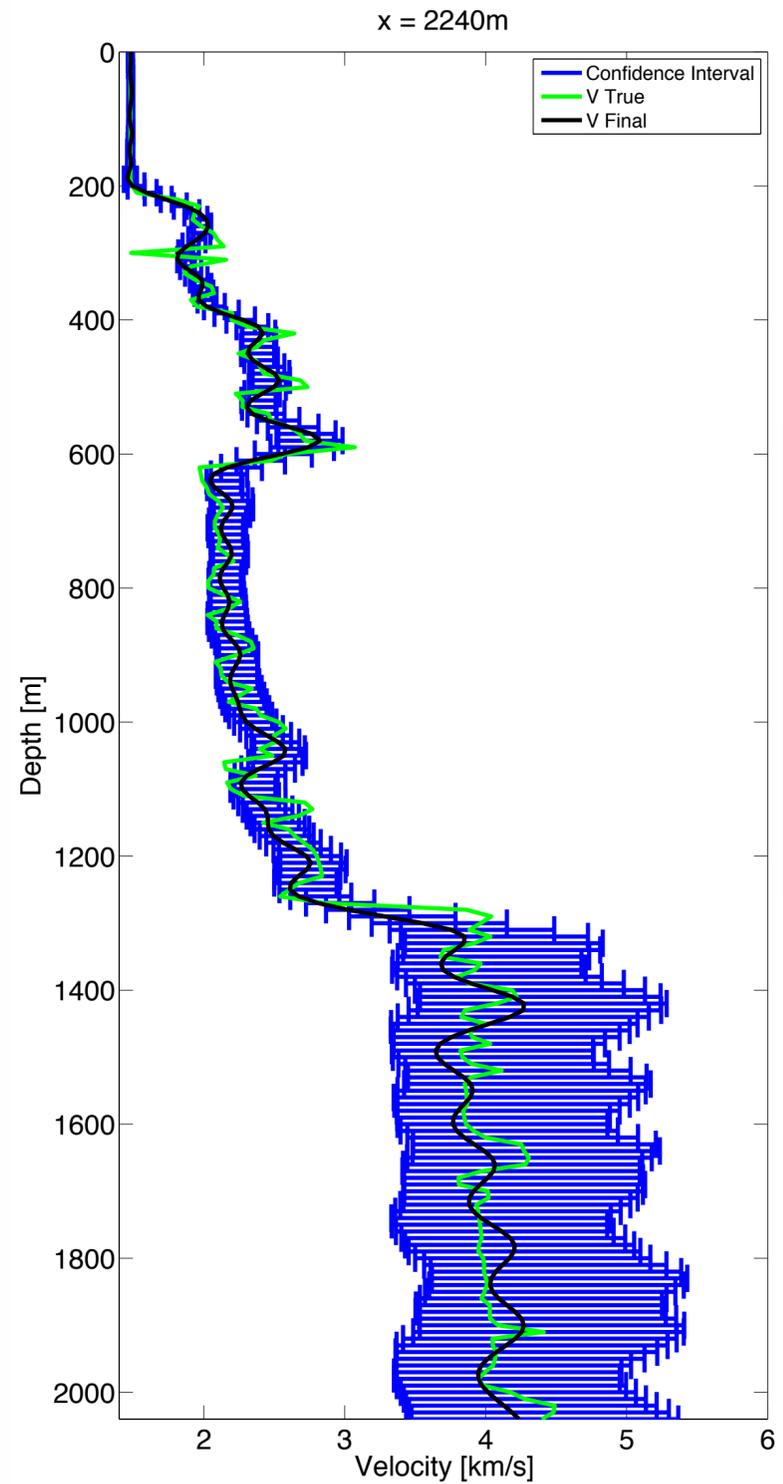


x = 1490m

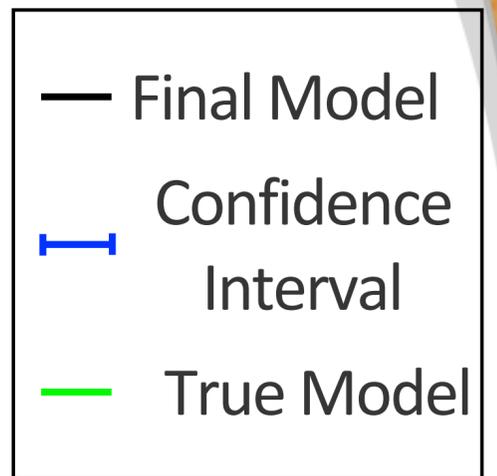


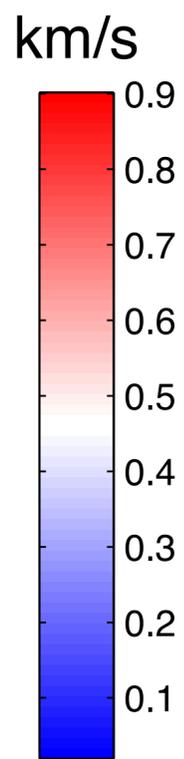
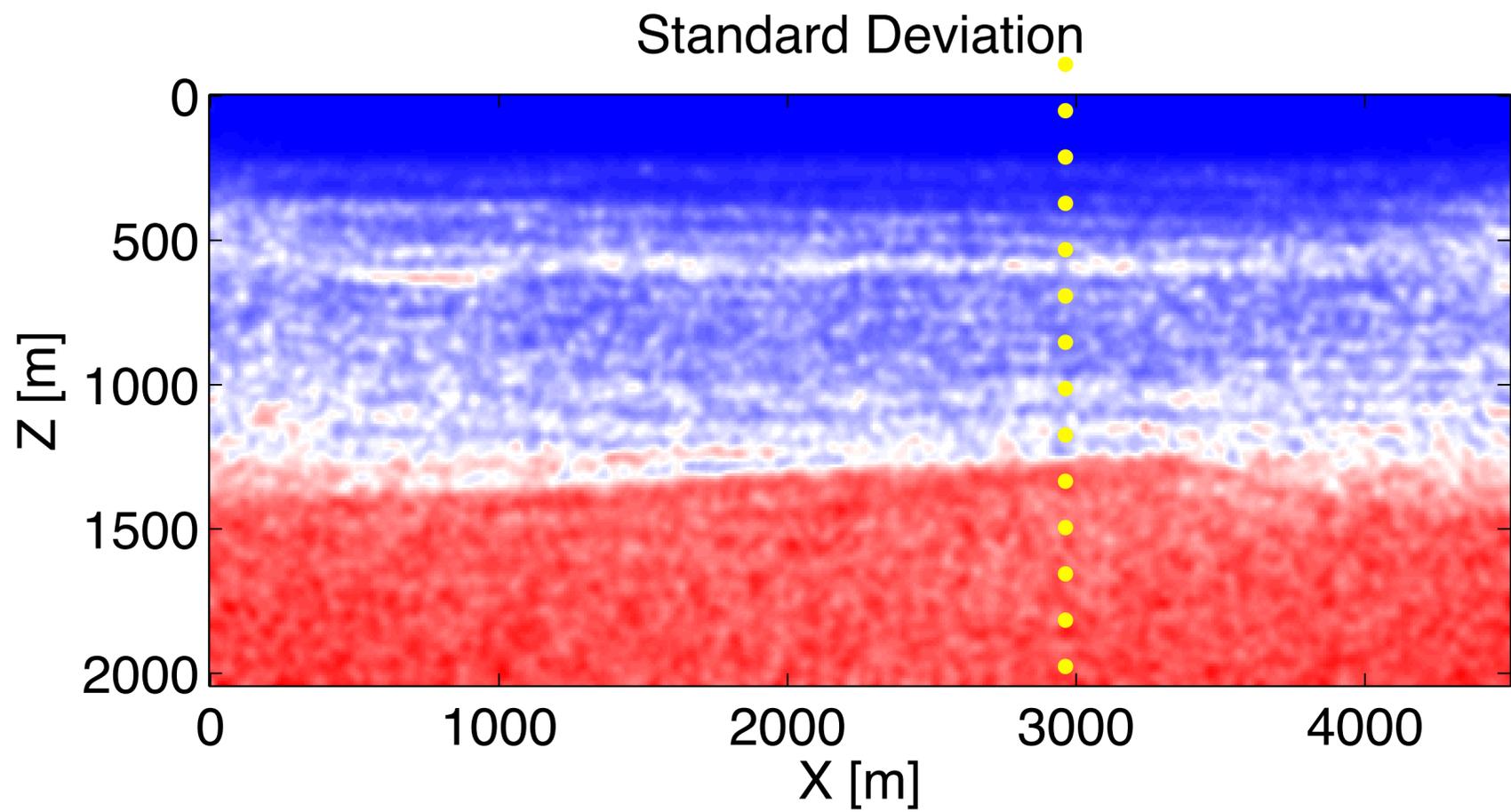


Standard deviation

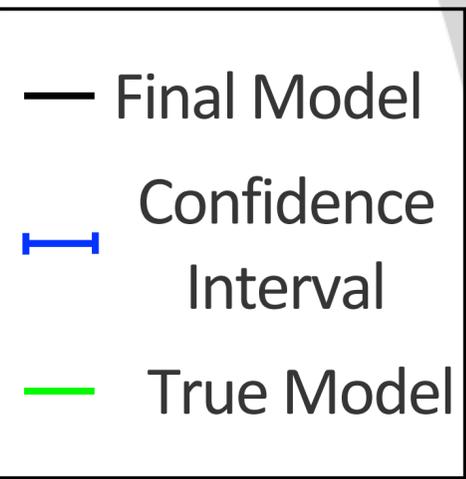
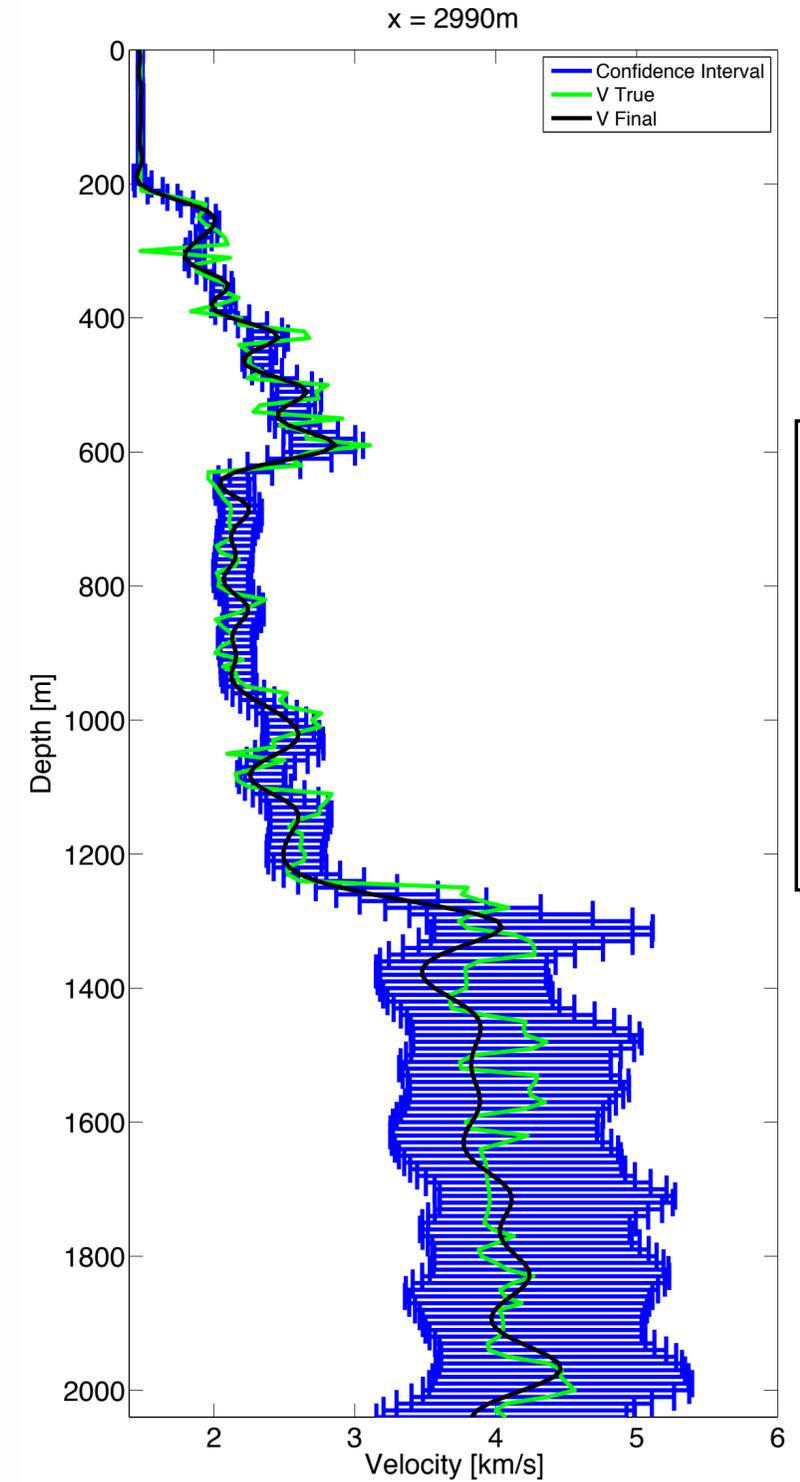


x = 2240m





Standard deviation



x = 2990m

## Conclusions

- Using the L-BFGS Hessian, we reduce the computational cost of estimating the Hessian.
- Using the randomized source sub-sampling method, we decrease the computational cost of the McMC.

## Future Work

- Use penalty method frame work to reset up the probability distribution function.
- Use different sampling method to analyze the posterior distribution.
- Uncertainty quantification for 3D FWI.

## Acknowledgements

Thank you for your attention !

<https://www.slim.eos.ubc.ca/>



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